



Combining fisheries surveys to inform marine species distribution modelling

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23

24

25 **Abstract** (228 words)

26 Ecosystem-scale examination of fish communities typically involves creating
27 spatiotemporally-explicit relative abundance distribution maps using data derived from
28 multiple fishery-independent surveys. However, survey sampling performance varies by
29 vessel and sampling gear, which may influence estimated species distribution patterns. Using
30 generalised additive mixed models, the effect of different gear-vessel combinations on
31 relative abundance estimates at length are investigated using European fisheries-independent
32 groundfish survey data. We constructed a modelling framework for evaluating relative
33 efficiency of multiple survey gear-vessel combinations and examined 19 disparate surveys
34 for 254 species-length combinations across the northeast Atlantic. Space-time variables
35 explained the majority of the variation in catches when combining data across different gears
36 or vessels for 181 of 254 species-length cases, indicating that for many species, models could
37 successfully characterize distribution patterns by combining data from disparate surveys.
38 Variables controlling for catch efficiency differences across gear-vessel combinations
39 explained substantial variation in catches for 127 of 254 species-length data sets. In such
40 cases, models that fail to control for gear efficiencies across surveys can mask changes in the
41 spatial distribution of species. Estimated relative differences in catch efficiencies grouped
42 strongly by gear type, but did not exhibit a clear pattern across species' functional forms,
43 suggesting difficulty in predicting the potential impact of gear efficiency differences when
44 combining data across surveys to assess species' distributions and highlighting the
45 importance of modelling approaches that can control for gear differences.

46

47 **Keywords**

48 Catchability; gear efficiency; fisheries independent assessment; Generalised Additive Mixed
49 Model (GAMM); survey standardisation; species distribution modelling

50 **1. Introduction**

51 As ecosystem-based management in the marine environment advances, fisheries policies
52 increasingly require consideration of both target and non-target species in assessing the state
53 of fisheries and fishing impacts on marine ecosystems (e.g. the European Union (EU) Marine
54 Strategy Framework Directive (MSFD; EC 2008; 2010; 2017), Common Fisheries Policy
55 (CFP; EC, 2013), United States Magnuson–Stevens Fishery Conservation and Management
56 Act (US, 1996; 2006), etc.). This transition to ecosystem-based management has led to a need
57 for greater understanding and detailed information on the distribution of a broad spectrum of
58 fish species across large spatial scales, such as large marine ecosystems or ecoregions (Kelley
59 & Sherman 2018).

60

61 Fisheries-independent groundfish surveys sample both commercial and non-target fish
62 species, often providing the only data source available to estimate relative abundances for
63 non-commercial species (Poos et al., 2013). These surveys tend to be discrete monitoring
64 programmes, operating at local scales usually associated with the exclusive economic zones
65 of countries managing the surveys. To obtain information on fish distributions at large marine
66 ecosystems scales, therefore, requires integration across national jurisdictional boundaries
67 and multiple disparate surveys that may differ in terms of spatial coverage, survey vessel,
68 season, types of fishing gear, and survey protocols. Amalgamating such data into a single
69 cohesive analysis is difficult because of potential differences in gear efficiency among
70 different length-classes and species of fish (Fraser et al., 2007; Walker et al., 2017), types of
71 survey gear, and vessels that vary in their fishing power (Dann et al., 2005).

72

73 Estimates of species' latent abundance, and hence species-at-length catchability coefficients,
74 are rarely available in fisheries survey data. In isolation, each individual survey provides

75 estimates of species' relative abundance at sampled locations and can provide assessment of
76 the spatial distribution of fish within the survey domain. Problems may arise, however, when
77 two or more surveys need to be combined to assess species' distributions. If gear efficiencies
78 vary between different surveys, then estimates of species relative abundance provided by
79 each survey may not be compatible. Failure to understand, or ignoring, how gear efficiency
80 differs between surveys may lead to incoherent abundance estimates when merging surveys
81 together to conduct assessments at large spatial scales. To perform such assessment,
82 therefore, requires quantification of gear efficiency for different species, different size classes
83 of fish, and different gears.

84

85 The traditional approach to estimating gear efficiency is through paired field experiments,
86 where two vessels fish side by side and compare catches (Somerton et al., 1999; Zhou et al.,
87 2014). Such experiments are costly to conduct and are generally implemented over limited
88 spatial and temporal scales. However, where different survey domains overlap spatially, there
89 may be opportunity to utilize species distribution modelling to complement, or even replace,
90 field-based estimation of gear efficiencies (e.g. Ono et al., 2018); thereby providing a
91 convenient framework for handling data from disparate surveys that can be regularly updated
92 as new survey data become available. Statistical modelling of species distributions from large
93 data sets is no longer limited by insufficient computing capacity. The use of such models
94 offers an opportunity of overcoming challenges in combining data across surveys with
95 varying gear efficiencies to enable extensive study of marine species distributions across
96 large spatial scales.

97

98 Here we build from previous gear efficiency modelling efforts (Walker et al., 2017, Zhou et
99 al., 2014) with an aim to advance the tools available for combining information across

100 disparate fisheries surveys towards informing the spatial ecology of marine species. The
101 spatial scale, the number of species assessed, the interaction between the gear-vessel
102 combinations, and the spatial and temporal variation inherent within European fisheries
103 surveys presents unique challenges requiring a new approach. Utilizing Generalised Additive
104 Mixed Models (GAMMs); we analyse the proportion of variance explained by the differences
105 in gear efficiency and the spatial–temporal variation in abundance of 135 species, in three
106 length categories, collected in the 19 northeast Atlantic groundfish surveys with 24 different
107 gear-vessel combinations. Here we focus on bottom trawl gears, namely otter trawls and
108 beam trawls, as others have previously focused on combining acoustic measurements with
109 habitat data to gain inference about the abundance of fish and infer on bottom trawl gear
110 efficiencies (Kotwicki et al., 2018). Three length categories were chosen to (1) capture the
111 main intra-specific length-related catchability differences described in previous studies
112 (Fraser et al., 2007; Walker et al., 2017), (2) broadly reflect trophic guilds in marine fish
113 communities (ICES, 2017), and (3) reflect the main size classes of fish either retained in
114 commercial trawls or that escape through the mesh (Piet et al., 2009). The 24 gear-vessel
115 combinations were chosen to best reflect the perceived differences in rigging and standard
116 operating procedures carried out by different countries in their national surveys (Table 1). By
117 understanding which species in our length categories are affected by variations among gears
118 and vessels, our primary goal is to develop a consistent approach for combining groundfish
119 surveys to facilitate marine ecosystem monitoring at large spatial scales. Using the GAMMs
120 to control for differences in gear efficiency among surveys, we also generate estimates of
121 spatial and temporal trends of relative abundance for species among different length
122 categories throughout the northeast Atlantic to inform marine fish community ecological
123 analyses (covering three ICES marine ecoregions/large marine ecosystems: Greater North
124 Sea, Celtic Seas, and Bay of Biscay and the Iberian Coastal; Spalding et al., 2007). Finally,

125 we conclude with a discussion of high priority information needs to further improve
126 understanding of gear efficiency within marine fisheries survey data.

127

128 **2. Methods**

129 *2.1 Fisheries Surveys*

130 Data for most European groundfish surveys are uploaded and maintained on the ICES
131 “Database of Trawl Surveys” (DATRAS). Data for surveys carried out in the Northeast
132 Atlantic were recently subjected to a quality assurance and quality audit (QAQA) process
133 (Moriarty et al., 2017; Greenstreet and Moriarty 2017a; 2017b; Moriarty et al., 2019), to
134 ensure their adequacy to meet monitoring and assessment requirements under the EU MSFD
135 (EC, 2008; 2010; 2017). These standard monitoring programme data products, along with
136 data for four Spanish surveys, which underwent the same QAQA process but were not fully
137 uploaded to DATRAS, were used in this study to obtain maximum spatial and temporal
138 coverage and include the widest possible range of survey types for modelling (Table 1). Each
139 survey data product includes the number of fish caught ($C_{i,s,l}$) of a species (s) at length (l),
140 for each trawl sample (i), along with the vessel and fishing gear (g), tow location, date,
141 depth and swept area (E). The fishing gear (g), included information from vessels that were
142 expected to fish differently based on their gear configuration information. For example, both
143 French and Irish vessels surveying in the Celtic Seas region use a GOV gear. However, the
144 French surveys use double sweeps, and the Irish surveys rotate between a standard GOV
145 survey gear (ICES 2015) and a double sweep with 16-inch bobbins, depending on the
146 substrate (Table 1). The fish abundance data were organized into three broad length
147 categories (lc), small unfished (<23cm), intermediate transition (23 - 35cm), and large fished
148 (>35cm). Groundfish surveys only record those species and lengths caught (i.e. presence only
149 data). Data rows for zero catches were added to the full data set where species at length were

150 not reported in any given sample. To ensure constant and equivalent distance units, survey
151 sample latitude - longitude coordinates were converted to eastings and northings (X, Y) using
152 R package “Rgdal” (Bivand et al., 2018). Date (t) was incremented in quarterly time bins
153 starting from quarter 4 (Oct – Dec) 2003, which was assigned time step $t = 1$, while the
154 quarter 1 (Jan – Mar) 2004 was assigned time step $t = 2$, and so on.

155

156 *2.2 Exploring Sources of Variation in Survey Abundance at Length Data*

157 Generalised Additive Mixed Models (GAMMs) were used to account for non-linear spatial
158 and temporal trends in fish density while simultaneously estimating gear efficiency using a
159 modelling framework adapted from Walker et al. (2017). Survey catches were modelled as
160 counts, with separate regressions for each species-length bin combination. Many species had
161 a preponderance of zero catches. Initial exploration casting GAMMs for all species within
162 Poisson, negative binomial, and zero-inflated Poisson frameworks showed that Poisson
163 models provided a poor fit and failed to accommodate over-dispersion in catch data. Negative
164 binomial and zero-inflated Poisson models showed similar fits for non-schooling species, but
165 schooling species violated the assumption of independence required by Poisson processes.
166 Consequently, we analyzed catches as Negative Binomially (NB) distributed GAMMs fit
167 using the “mgcv” package (Wood 2004; 2011) in the R statistical programming environment
168 (R Core Team 2017). The full model for a given species and length category catch data set
169 had the form:

$$170 C_i \sim NB(\mu_i, k)$$

$$171 \text{ with } E[C_i] = \mu_i = e^{\log(E_i) + s(X_i, Y_i, t_i) + zg_{(i)}} \quad 1,$$

172 where C_i is the number of fish of a given species in a given length category caught in the i^{th}
173 sample (fishing event), k is the negative binomial shape parameter representing the degree of

174 overdispersion, $\log(E_i)$ is the log of swept area for fishing event i which was included as an
175 offset to account for varying fishing effort among trips, $s(X_i, Y_i, t_i)$ denotes a multivariate
176 smoothing function to represent spatio-temporal trends in catch data, and $z_{g(i)}$ are i.i.d.
177 normally distributed random effects for gear-vessel combinations associated with fishing
178 events. The space-time smoothing model component, $s(X_i, Y_i, t_i)$, was specified as a tensor
179 product smoother for which the associated basis functions were cast as cubic splines with
180 shrinkage (i.e., $te(X_i, Y_i, t_i, bs = "cs")$ in mgcv formulaic notation), a formulation which can
181 accommodate data on different scales (Wood 2004; 2011). Gear-vessel combination was
182 treated as a random effect, as opposed to a fixed effect, because variation among catch
183 efficiencies is the primary feature of interest, and because this approach also aids in model
184 convergence by reducing the number of fitted parameters. The spatiotemporal smoother
185 describes the underlying estimated distribution of species across space and time; whereas the
186 random effect controls for variation among gear efficiency when combining disparate survey
187 data sets. To facilitate model convergence, we excluded data on species-at-length for which
188 any given length category was sampled by fewer than two gear-vessel combinations or was
189 sampled fewer than 100 times. The full model was compared to a reduced model that
190 included space-time covariates, but which did not account for the effect of gear-vessel
191 combinations (i.e. the gear-vessel combination random effect was dropped) in order to assess
192 the impact on species distribution modelling inference when gear is ignored. Comparisons of
193 full and reduced model fits were assessed using Akaike's information criterion (AIC). The
194 full model was further assessed for reliability using visual tests and a chi squared goodness of
195 fit test. To substantiate that our GAMM models can effectively differentiate between the
196 random gear-vessel effects and the spatial and temporal variation in the abundance of
197 demersal fish in the north east Atlantic region, we performed a simulation-estimation
198 experiment (Supplemental Material S2).

199

200 *2.3 Interpretation of models*

201 To interpret the importance of gear efficiency versus spatiotemporal distribution patterns in
202 explaining variation in survey data, we utilized variance components analysis. This analysis
203 partitions total variation in the fitted data among the three modelled components: gear
204 efficiency, spatiotemporal distribution, or unexplained residual variation. Accordingly, when
205 the gear component constitutes the preponderance of model variation for a given species and/
206 or length category, we conclude that gear efficiency varies widely across gears and surveys.
207 In contrast, when location and time make up the majority of model variability for a given
208 species, we conclude that catches are more strongly influenced by the ecology of the fish,
209 rather than the differences in gear efficiency.

210

211 A non-metric multidimensional scaling (nMDS) unconstrained ordination technique using
212 Euclidean distances was employed to explore how each species within the assemblages
213 varied with estimates of gear efficiencies among gear-vessel coefficients and length classes
214 from our models. Species were grouped by taxonomic order as a proxy for functional forms
215 to examine if there was a pattern in estimates of gear efficiencies in species groups with
216 similar morphological or ecological attributes. The gear-vessel coefficients were conditioned
217 into a matrix, where the Scottish vessel with a GOV gear type was used as a reference gear,
218 and the difference was calculated for each other gear-vessel combination. Permutational
219 multivariate analysis of variance (PERMANOVA) was used to test the differences between
220 the gear-vessel coefficients derived for each species in each length class from our full models
221 for similar gear types. A clustering criterion that minimizes the amount of variance within in
222 the gear-vessel groups was implemented (Ward, 1963). Euclidean distance was used and the

223 *p*-value was set to 0.05. The nMDS and PERMANOVA routines were implemented in R (R
224 Core Team 2017) using the “vegan” package (Oksanen et al., 2017).

225

226 **3. Results**

227 Data for 135 fish species were available from otter trawl surveys across the northeast
228 Atlantic, whereas beam trawl surveys operate in a much more limited area within the North
229 and Irish Seas (Figure 1). The surveys carried out in the Irish Sea have the highest degree of
230 spatial and temporal overlap, whereas survey overlap is more limited in the Bay of Biscay
231 and Iberian Coast region (Figure 1).

232

233 Two hundred and fifty-four full GAMMs were fit to 132 species in up to three length
234 categories (Figure 2). For fishes in the smallest size class (<23cm), the full model was fit to
235 109 species, and 23 species had insufficient data based on the criteria described in Methods
236 (Section 2.2). For fishes in the intermediate transition category (23-35cm), the full model was
237 fit to 85 species, and 47 species had insufficient data. For the largest size class (>35cm), the
238 full model was fit to 60 species, and 72 species had insufficient data.

239

240 In 39/254 models, the unexplained variance was greater than the explained variance (Figure
241 2). In 237/254 of the species-length combinations, the full model, which controlled for
242 differences in gear-vessel combinations, improved the deviance explained over the reduced
243 model (Table S1.1). 250/254 full models had a lower AIC score than the reduced model. In
244 the cases where the full estimates did not improve inference, the differences in the amounts of
245 deviance explained and the AIC scores between the full and reduced models were small
246 (Table S1.1).

247

248 In 215/254 full models, over 50% of the variation in the data can be explained, suggesting
249 that this framework is an effective way of calculating variance in latent species abundance
250 over a large spatial scale. In 181/254 full models, location (X, Y) and time (t) components of
251 the model explained over 50% of the variation in the data, suggesting that catch rates are
252 strongly driven by the ecology of the fish, while the random effect of fishing gear on a given
253 vessel(g) at a given length category (l) generally plays a smaller role in explaining variance.
254 Indeed, in 51 of these 181 models, the overall variance explained is >50%, but the variance
255 explained by gear is <1%. As an example, for common dab (*Limanda limanda*) in the <23cm
256 length class, the random effect of fishing gear on a given vessel (g) explains 0.007% of the
257 variance, while location (X, Y) and time (t) components explained 62.2% of the variance
258 (Figure 3a/b). In this case, the reduced model, where location (X, Y) and time (t) components
259 explained 61.1% of the variance, performed similarly to the full model (Supplemental
260 Material 1 Table S1.1).

261

262 In 37/254 full models, the overall variance explained is >50%, and the gear component
263 explains between 1% and 5% of the variation, suggesting that gear efficiency varies across
264 gears and vessel combinations but has relatively little influence on catch performance. For
265 example, for the thorny skate (*Amblyraja radiata*) in the 23-35cm length class, the random
266 effect of fishing gear (g) explained 3.7% of the variance, while location (X, Y) and time
267 (t) components of the full model explained 68.7% of the variance. While the estimated
268 variance component for gear effects was smaller than the space-time components, the effect
269 of fishing gear can be seen in the difference in spatial pattern between the full and reduced
270 models (Figure 3d).

271

272 In 127/254 full models the overall variance explained is >50%, and the gear component
273 explains more than 5% of the variation, suggesting that gear efficiency for these species-at-
274 length varies substantially across gear and vessel combinations. For example, for sole (*Solea*
275 *solea*) in the 23-35cm length class, the random effect of fishing gear (*g*) explained 8.6% of
276 the variance, while location (*X, Y*) and time (*t*) components of the full model explained
277 46.5% of the variance in the data (Figure 3e/f). In this case, the output of the full model
278 highlights the importance of understanding the effect of fishing gear in assessing the
279 distribution of this species.

280

281 To assess the difference in inference gleaned from the full and reduced models, we further
282 explored the spatial-temporal pattern of sole (*Solea solea*) in the 23-35cm category. While the
283 general pattern is similar in the full and reduced models (Figure 4), the reduced model
284 suggests the presence of intermediate-sized sole off of the coasts of Spain and Portugal;
285 whereas the full model suggests that there are no intermediate-sized sole in these areas. When
286 examined more specifically, we see that for the entire area, the sole data is 88% zero values,
287 but for the southern part of the study area, where Spain and Portugal survey, the sole data is
288 96.5% zero values. Consequently, we can conclude that the reduced model is likely to
289 overestimate the abundance in this area, and that this overestimation is likely an artefact of
290 not accounting for gear.

291

292 Aggregating over the entire distribution of sole, there is a steadier rate of movement in the
293 centre of mass in the population estimated from the full model, while the movement in the
294 centre of mass in the population estimated from the reduced model is more variable (Figure
295 5a). The centre of mass metric highlights the eastward movement in the population in the full
296 model, which is not the case in the reduced model (Figure 5b). The inference from the

297 simulations suggests that the full model should be more capable of capturing the direction of
298 movement than the reduced model ((Supplemental 2, Figure S2.4).

299

300 Unsurprisingly, nMDS highlights that the estimated gear coefficients vary considerably by
301 gear types (Figure 6a; PERMANOVA test for differences in gears: $F = 2.36, R^2 = 0.18, p -$
302 $value = 0.001$). However, gear coefficients are largely consistent within gear type,
303 indicating stable catch efficiencies within gear types regardless of the survey country of
304 origin or vessel. The GOV, beam trawls, and baca trawls gear-vessels tended to group most
305 closely in their estimated gear coefficients, whereas other gears tended to differ more widely.
306 The GOV has the highest level of variance and is the most widely used gear within the
307 region. The beam trawl surveys have a high level of spatial overlap with the surveys that use
308 the GOV gear in the North Sea and the rockhopper trawl in the Irish Seas. The baca trawl has
309 very limited spatial overlap with other gears as it is used exclusively by the Spanish in the
310 Bay of Biscay and Iberian Coast region. There is no clear pattern emerging in the estimated
311 relative difference in catch efficiencies across species functional form (Figure 6b).

312

313 **4. Discussion**

314 Understanding how gear efficiency impacts fishery independent survey sampling is required
315 for robust multi-survey species distribution modelling of both commercial and non-
316 commercial species and is a key factor in determining absolute abundance estimates for
317 commercial stocks (Kasatkina & Ivanova, 2009; Maunder & Piner, 2014). The aim of the
318 analyses presented here is to provide an overall understanding how species are affected by the
319 rigging of individual vessels to guide future ecosystem-scale species distribution modelling
320 and examinations of fish communities. Our models support the derivation of relative species
321 abundance estimates, and they provide information on gear efficiency of 24 gear-vessel

322 combinations seasonally for three length groups chosen to reflect the main intra-specific
323 length-related differences described in previous catchability studies (Fraser et al., 2007) in
324 this region. This provides a modelling workflow to combine data across surveys that controls
325 for potential gear-vessel-specific differences in catchability. The flexible framework
326 provided here may be adapted to the end users' needs; for example, different length
327 categories may be applied to answer specific ecological questions. We caution; however, that
328 the gear efficiency coefficients used in this analysis were estimated using a 10-year historical
329 time span and are only valid under the conditions for which they are calculated. As such, any
330 efforts to employ them for correcting individual survey-species catches need take this into
331 account (Arreguín-Sánchez, 1996).

332 In 15% (39/254) of models, the unexplained variance is higher than the explained variance
333 (Figure 2). Given that it is unlikely for a species to be randomly distributed in space and time,
334 this high unexplained variance is likely due to the rareness of the species within a given
335 length category (i.e. there are not enough samples to describe the latent species distribution).
336 Species that are rarely caught may not be rare in the environment, but instead may be
337 particularly poorly sampled (i.e. low gear efficiency) in the survey trawl gear. Sampling of
338 fish in the marine environment by fishing gear is known to be imperfect (Fraser et al., 2007,
339 Zhou et al., 2014, Walker et al., 2017). This means additional considerations may need to be
340 addressed during sampling and data analysis, such as joint dynamic species distribution
341 modelling (Thorson et al., 2016). Reliable inference depends on sampling methods that
342 produce reasonable odds of detection given presence, where no estimator will be particularly
343 helpful when applied to data on populations or species that are “invisible” to collection gear
344 (MacKenzie et al., 2006).

345

346 The estimated variance components from our models show that in 35% of cases (88/254), the
347 location and time components explained most of the variation in the data, while the gear
348 component explained relatively little variation ($\leq 5\%$; Figure 2). This suggests that in such
349 circumstances, the spatial-temporal distribution of these species can be estimated using
350 combined survey data. Where the modelled gear component is especially small, particularly
351 in relation to the location and time component, use of raw survey catch data from multiple
352 surveys provides a reasonably accurate representation of temporal and spatial variation in
353 species' abundances (by length category) at large spatial scales. The common dab (Figure
354 3a/b.), highlights a circumstance in which little variance can be attributed to gear effects, and
355 we see a consequent small difference in inference in the temporal and spatial trends between
356 the full and reduced models. The variance explained by the gear is $<1\%$ while the spatial and
357 temporal components explain 62.2% of the variance. Thus, this species (by length category)
358 abundance appears to be less impacted by the effects of gear as the catch rates are likely
359 driven by the ecology of the fish. The variation that is attributable to gear effects is smaller
360 than that attributed to space and time in most of our GAMM models, but the nature of the
361 gear effects are not randomly distributed throughout the study area or throughout the year.
362 They are instead systematically distributed by seasonal surveys. This regularity in the
363 differences may impact species distribution inference at large scales. Simulations (S2a) for
364 species demonstrating substantial movements in distribution attributed 5.7 % of model
365 variance to gear, even when no gear effect was included. This suggests that some of the
366 variance associated with location and time may be attributed to gear, but inferences from full
367 and reduced models were similar. Conversely, when there is a strong gear effect (S2b) then
368 the full model improves inference of abundance estimates and direction of population centre
369 of mass movements over the reduced model (Supplemental Material 2).

370

371 Not accounting for gear may lead to incorrect estimates of relative abundance or species'
372 distributions. Data analysed here suggest that gear effects on catches across disparate surveys
373 are not uncommon, whereby in half of our full models (127/254), the gear component
374 explained more than 5% of the total variation in survey catches, while overall variance
375 explained is >50%. Our examination of the distribution of sole provides demonstration of the
376 potential importance of controlling for gear effects when attempting to combine data across
377 surveys for some species. The variance explained by gear in this case was 8.6%, while the
378 spatial and temporal components of the model accounted for 46.5% of the variance.
379 Consequently, we found substantial differences in relative abundance trends between models
380 which control for gear effects compared to reduced models which ignore gear effects in
381 combining data across surveys (Figures 3d/e, 4, 5). Importantly, failure to control for gear
382 differences across surveys for this species would mask differences in the spatial distribution
383 of the stock across commercial fishing areas, as well as mask ecosystem-scale population
384 shifts to the east (Figure 5). It may be valid to pool across surveys in assessing species
385 distributions for many species-size combinations; however, there are differences evident
386 across gear types and it is not clear a priori for which species gear differences matter (Figure
387 6b). Thus, a sensible workflow when combining data across surveys may be to implement
388 models that control for gear type as demonstrated here and then subsequently evaluate
389 whether gear differences account for a substantial portion of the variation in catches.

390

391 Northeast Atlantic waters are currently surveyed by 12 countries carrying out 19 different
392 surveys designed with individual goals and objectives and using different vessels and a
393 variety of gears (Table 1). ICES facilitates survey coordination and collaboration through
394 working groups to make the surveys as comparable as possible. The North Sea bottom trawl
395 surveys have led the way in terms of minimising gear efficiency issues caused by differences

396 in vessels and by ensuring survey overlap and similarity among gears (ICES, 2015). There is
397 a large body of work ongoing in ICES survey groups (e.g. WGBEAM, International Bottom
398 Trawl Survey Working Group; IBTSWG) to minimise survey variability; however, assessing
399 relative gear efficiency at the scale examined here highlights the need for comparative
400 experiments to help achieve a more coherent understanding of gear efficiency within fisheries
401 independent survey data. This is particularly relevant in the Bay of Biscay, where
402 overlapping or paired tows between the Spanish Baca Trawl and Portuguese Norwegian
403 Campelen Trawl and the Spanish Baca Trawl and French Grande Overture Vertical Trawl
404 would help to improve inferences of species relative abundance obtained from these different
405 gears (Figure 6a). Analyses herein provide further understanding of the differences in gear
406 efficiency between trawl gears used by different surveys for species sampled across the
407 northeast Atlantic.

408

409 Information on the abundance and distribution of organisms is a fundamental knowledge
410 need for fisheries management. Data on predator and prey abundances by different age and
411 size classes can inform species status assessments as well as provide information on the
412 interactions among species and size classes, providing understanding about the impact of
413 fishing on fish communities (Fraser et al., 2007; e.g. Large Fish Indicator). This study
414 provides an approach to facilitate comparability between catches from different surveys and
415 gears, providing a framework to assist in integrating data across countries, regions, and
416 sampling programs towards maximizing the use of available information to inform species'
417 abundance and spatial distribution assessments.

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424

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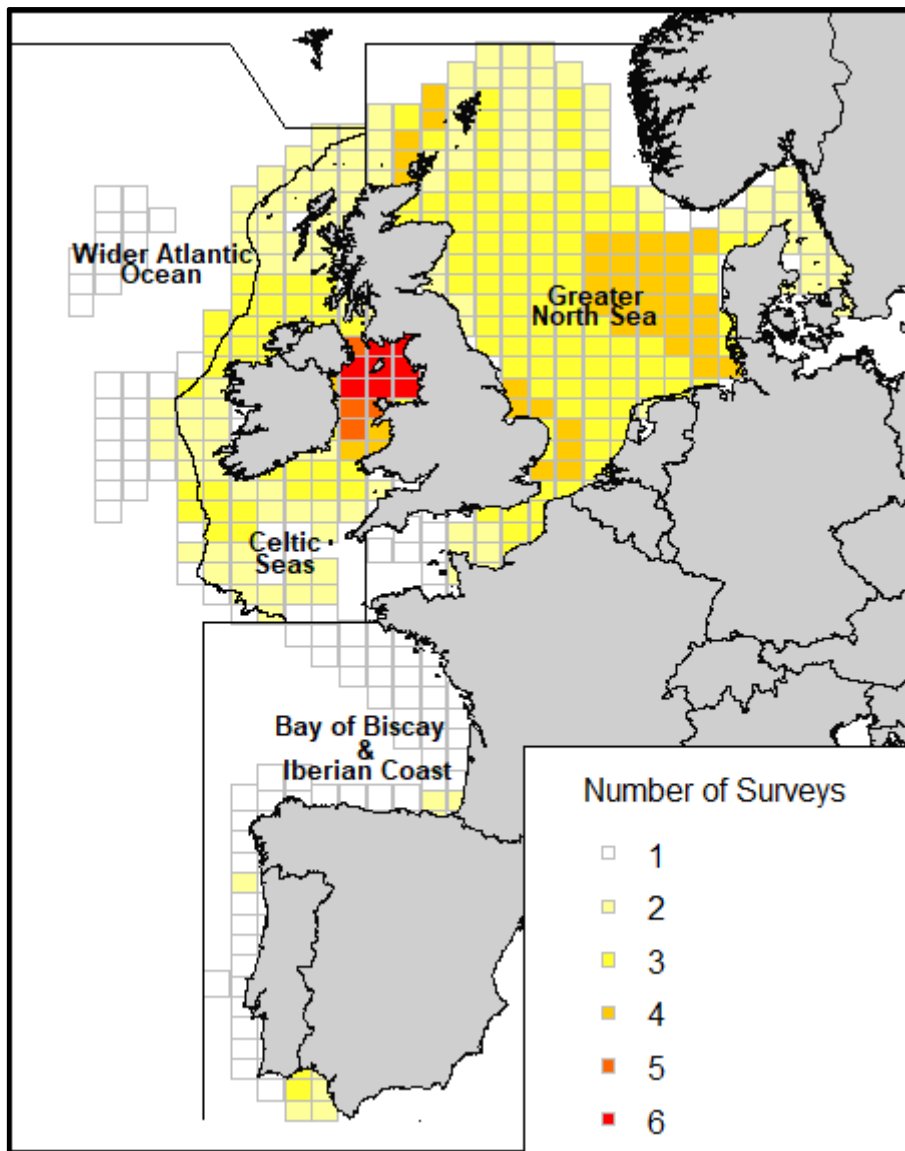
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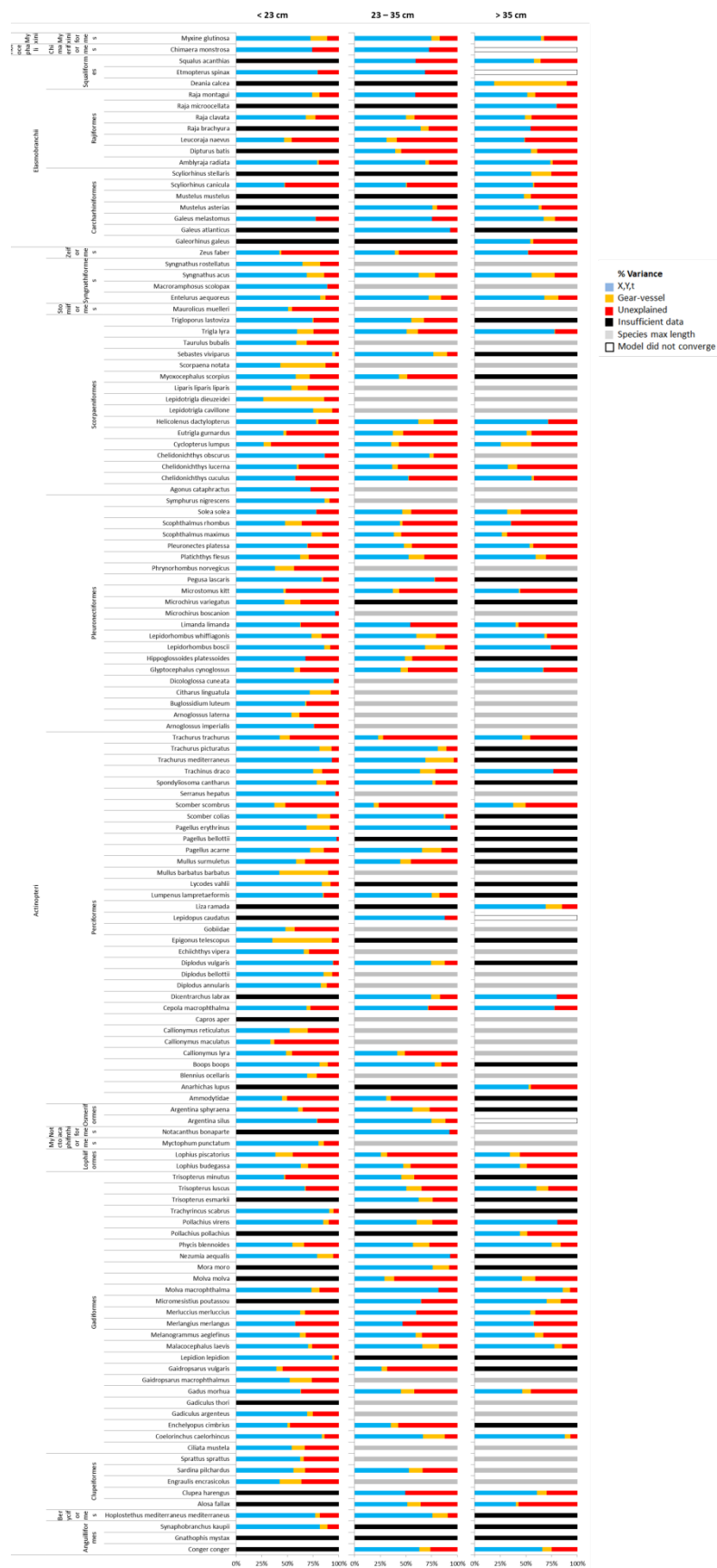
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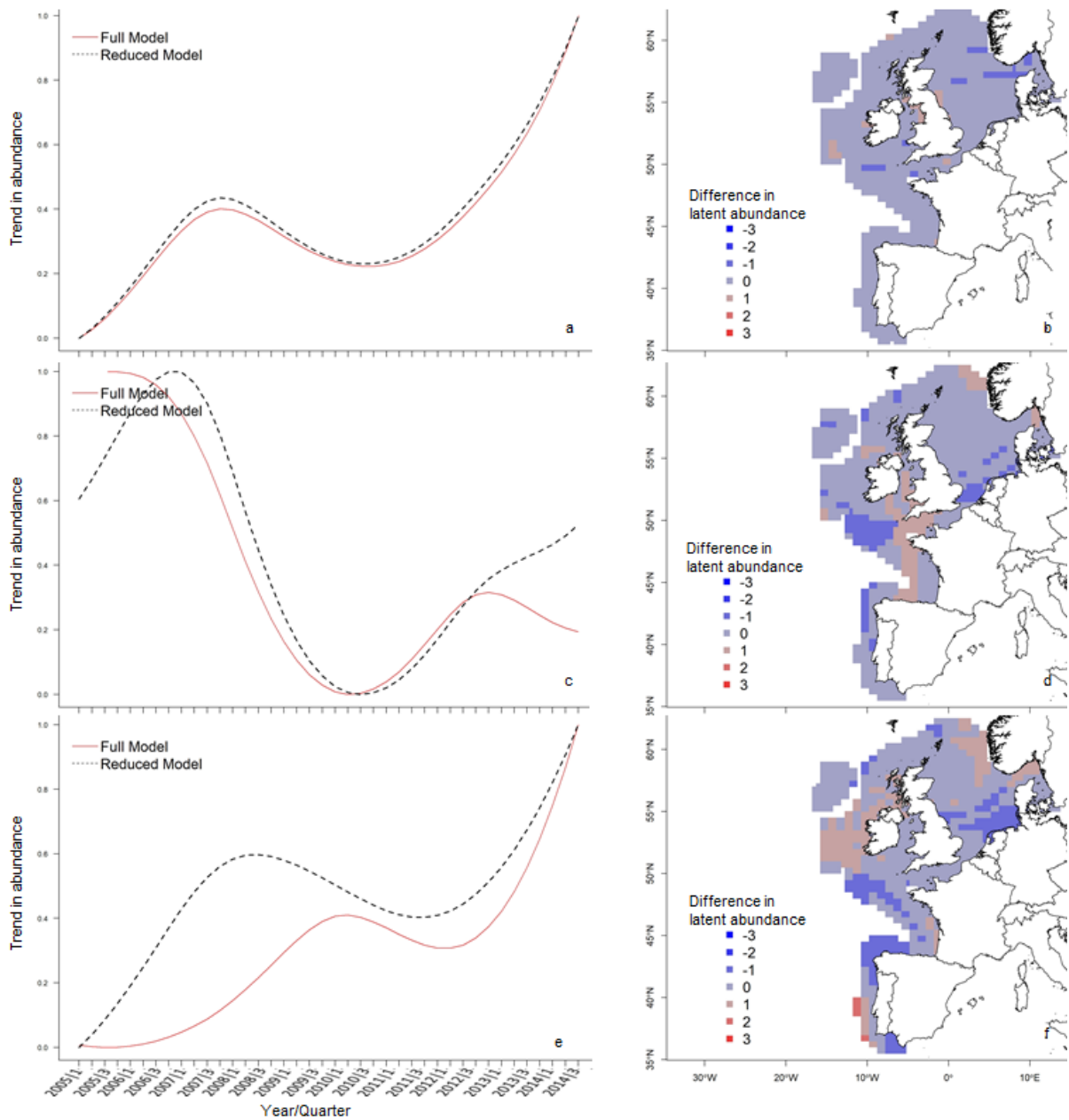
535
 536 **Figure 1:** Fisheries independent survey coverage across the northeast Atlantic. Thick black
 537 line shows Oslo/Paris convention (OSPAR) boundaries. Number of surveys operating in each
 538 ICES statistical rectangle is depicted by a different colour. See Table 1 for list of surveys.



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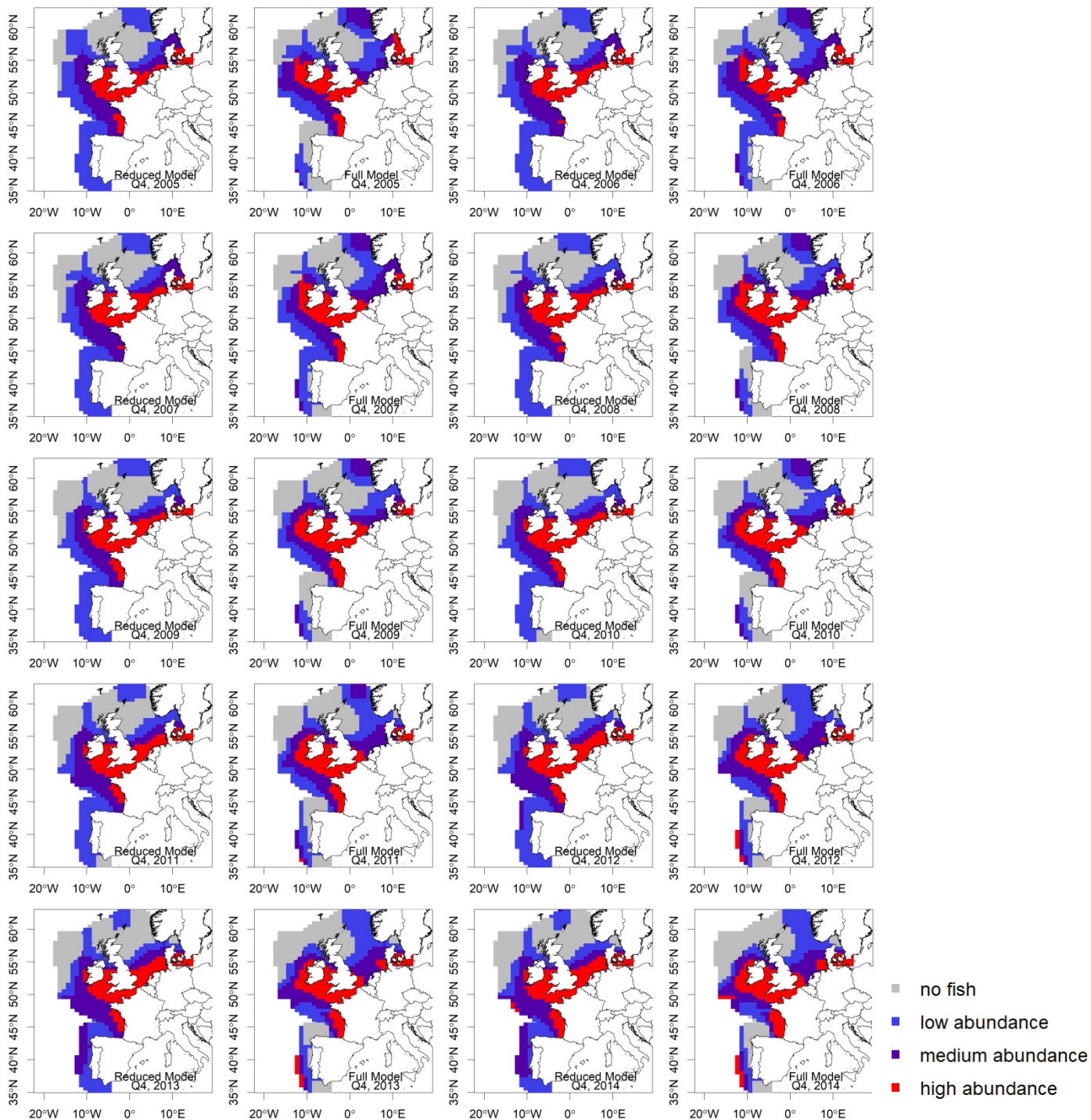
Figure 2: Summary of the proportion of variance explained from full model components for each length category (<23cm; 23 - 35cm and >35cm) and species, grouped in taxonomic order. X, Y and time (t) variance components are represented by blue bars, gear-vessel

543 components by orange bars, and unexplained variance by red bars. Black bars indicate
 544 insufficient data to fit a model for a given species-size combination, and white bars indicate
 545 model convergence failed. Finally, grey bars indicate a given length size bin is larger than
 546 the maximum observed length of a species.

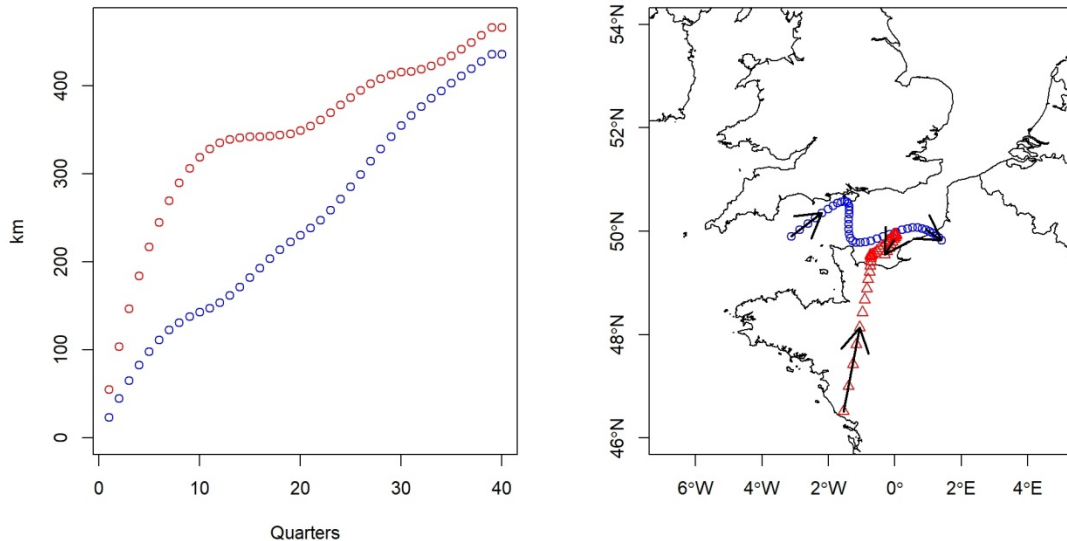


547
 548 *Figure 3: Top row (a,b): Common dab (Limanda limanda) < 23cm, highlighting an example*
 549 *of a species where the reduced and full model perform similarly as the variance explained by*
 550 *gear is very small (0.007%). Middle row (c,d) Thorny skate (Amblyraja radiata) 23 – 35cm,*
 551 *highlighting an example of a species with between 1-5% variance explained by gear. Bottom*
 552 *row (e,f) Sole (Solea solea) 23-35cm, highlighting an example of a species with >5%*
 553 *variance explained by gear. Left column (a,c,e): Estimated domain-wide species' abundance*
 554 *trends for the full model which controls for gear differences across surveys, versus the*
 555 *reduced model which does not control for gears. A large discrepancy between the curves*
 556 *indicates gear differences across surveys may impact inference about species' abundance*
 557 *and distributions. Right column (b,d,f): Differences in predicted species' relative mean*

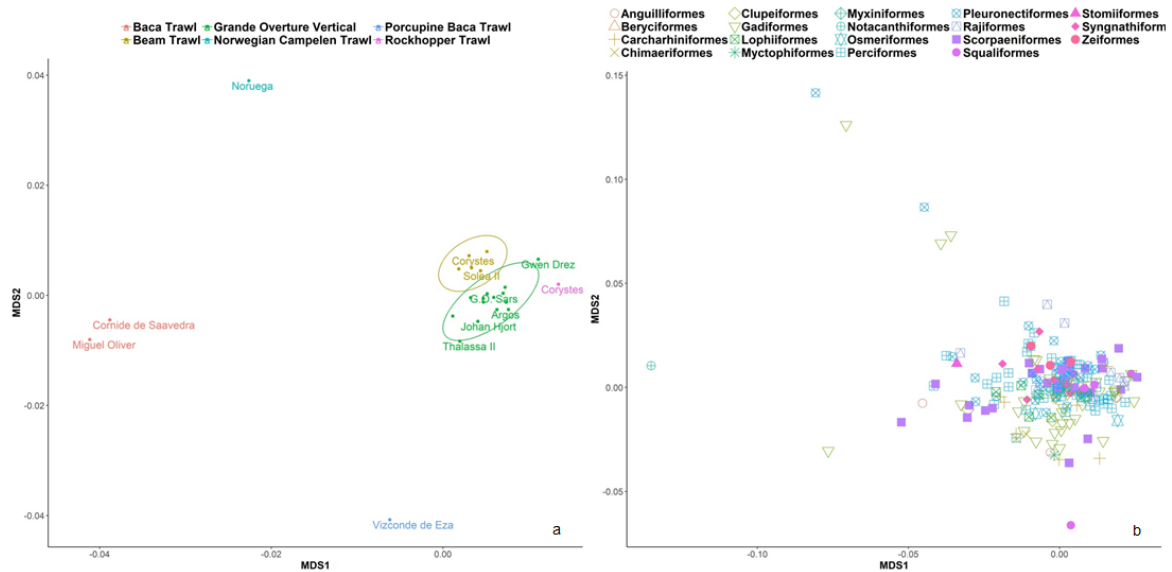
558 abundance between the full and reduced models. Dark colours represent large discrepancies
 559 between the models, indicating differences in gears across surveys may influence estimated
 560 species' distributions if not accounted for.
 561



562
 563
 564 *Figure 4. Spatial-temporal pattern in quarter 4 (Oct.-Dec.) for each year of sole (Solea*
 565 *solea) 23-35cm from the reduced model on the left and the full model on the right.*
 566 *Abundance is depicted as “low” in the 1st-2nd quantile, “medium” in the 2nd – 3rd*
 567 *and high in the 3rd-4th quantile.*



568
 569 *Figure 5. Summary of difference in inference from the spatial-temporal pattern of sole*
 570 *(Solea solea) 23-35cm from the full and reduced models. (a.) Cumulative movement from the*
 571 *centre of mass from the start of the time series for the full model (blue circles) and the*
 572 *reduced model (red triangles). (b.) Centre of mass of the abundance of the fish from the full*
 573 *model (blue circles) and the reduced model (red triangles).*
 574



575
 576 *Figure 6. Non-metric multidimensional scaling (nMDS) plots describing how the gear-vessel*
 577 *coefficients varied by survey or by taxonomic grouping (Stress = 0.102). (a) Gear coefficients*
 578 *grouped by trawl type (colours) and survey research vessel name (labels). Points more*
 579 *closely situated are more similar in terms of their gear-vessel coefficients. Ellipses indicate*
 580 *the 95% confidence intervals for clusters of each gear type. (b) Gear coefficients grouped by*
 581 *taxonomic order as a proxy for species' functional form.*

582

583

| Survey Acronym | DATRAS Acronym | Subregion | Country | Start Year | End Year | Quarter | Vessels | Gear Type | Mesh size (mm) | Haul Duration (min) $\bar{x} \pm s$ | Distance Towed (km) $\bar{x} \pm s$ | Wing Swept Area (km ²) $\bar{x} \pm s$ | Data Source | DOI |
|----------------|----------------|-------------------|---------------|------------|-----------|---------|-------------------------------|-------------|-------------------|--|--|---|-------------|----------------|
| GNSIntOT1 | IBTS | Greater North Sea | International | 1983 | 2017 | 1 | Multiple Ships | Otter (GOV) | | | | | DATRAS | 10.7489/1922-1 |
| | | | | | | | G.O.Sars | | 20 | 29±4 | 3.4±0.6 | 0.06±0.01 | | |
| | | | | | | | Argos | | 20 | 30±4 | 3.6±0.5 | 0.07±0.01 | | |
| | | | | | | | Dana | | 20 | 30±2 | 3.6±0.3 | 0.07±0.01 | | |
| | | | | | | | Dana (Sweden) | | 20 | 30±2 | 3.4±0.3 | 0.07±0.01 | | |
| | | | | | | | CEFAS Endeavour (Netherlands) | | 20 | 29±3 | 3.6±0.4 | 0.0±0.017 | | |
| | | | | | | | Haakon Mosby | | 20 | 29±3 | 2.9±0.3 | 0.06±0.1 | | |
| | | | | | | | Mimer | | 20 | 29±3 | 3.3±0.3 | 0.06±0.01 | | |
| | | | | | | | Scotia III | | 20 | 31±6 | 3.6±0.7 | 0.07±0.01 | | |
| | | | | | | | Thalassa II | | 20 | 30±1 | 3.6±0.4 | 0.06±0.01 | | |
| GNSIntOT3 | IBTS | Greater North Sea | International | 1998 | 2016 | 3 | Multiple Ships | Otter (GOV) | | | | | DATRAS | 10.7489/1923-1 |
| | | | | | | | Argos | | 20 | 30±1 | 3.5±0.2 | 0.07±0.01 | | |
| | | | | | | | Dana | | 20 | 29±4 | 3.6±0.5 | 0.07±0.01 | | |
| | | | | | | | Dana (Sweden) | | 20 | 30±1 | 3.4±0.2 | 0.07±0.01 | | |
| | | | | | | | CEFAS Endeavour | | 20 | 29±3 | 3.5±0.4 | 0.07±0.01 | | |
| | | | | | | | Haakon Mosby | | 20 | 27±5 | 3.2±0.6 | 0.07±0.01 | | |
| | | | | | | | Johan Hjort | | 20 | 27±6 | 3±0.8 | 0.06±0.02 | | |
| | | | | | | | Scotia III | | 20 | 29±4 | 3.3±0.5 | 0.06±0.01 | | |
| | | | | | | | Walther Herwig III | | 20 | 29±4 | 3.6±0.7 | 0.07±0.01 | | |
| | | | | | | | GNSFraOT4 | FR CGFS | Greater North Sea | France | 1988 | 2016 | | |
| Thalassa II | | 20 | 29±2 | 3.4±0.3 | 0.05±0.01 | | | | | | | | | |
| | | 20 | 29±3 | 2.9±0.5 | 0.03±0.01 | | | | | | | | | |
| CSScoOT1 | SWC-IBTS | Celtic Sea | Scotland | 1985 | 2016 | 1 | Scotia II | Otter (GOV) | 20 | 56±10 | 7.4±1.7 | 0.15±0.03 | DATRAS | 10.7489/1957-1 |
| | | | | | | | Scotia III | | 20 | 30±6 | 3.4±0.7 | 0.07±0.01 | | |
| CSScoOT4 | SWC-IBTS | Celtic Sea | Scotland | 1997 | 2016 | 4 | Scotia II | Otter (GOV) | 20 | 56±10 | 6.6±1.7 | 0.13±0.03 | DATRAS | 10.7489/1924-1 |
| | | | | | | | Scotia III | | 20 | 29±3 | 3.4±0.4 | 0.06±0.01 | | |
| CSreOT4 | IE-IGFS | Celtic Sea | Ireland | 2003 | 2016 | 4 | Celtic Explorer | Otter (GOV) | 20 | 30 ± 2 | 3.6±0.3 | 0.07±0.01 | DATRAS | 10.7489/1925-1 |
| CSNlrOT1 | NIGFS | Celtic Sea | Northern | 1992 | 2016 | 1 | Corystes, Lough | Otter (ROT) | | | | | DATRAS | 10.7489/1961- |

| | | | | | | | | | | | | | | | |
|---------------|----------|---------------------------------|------------------|------|------|---|-----------------------|---------------|----|--------|----------|------------|-------------------------------|----------------------|---|
| | | | Ireland | | | | Foyle | | | | | | | | 1 |
| | | | | | | | Corystes | | 20 | 58±9 | 5.3±0.9 | 0.08±0.01 | | | |
| | | | | | | | Lough Foyle | | 20 | 59±6 | 5.±0.6 | 0.08±0.01 | | | |
| CSNIrOT4 | NIGFS | Celtic Sea | Northern Ireland | 1992 | 2016 | 4 | Corystes, Lough Foyle | Otter (ROT) | | | | | NDB (92-07) DATRAS (08-15) | 10.7489/1962-1 | |
| | | | | | | | Corystes | | 20 | 19±1 | 1.9±0.01 | 0.03±0.02 | | | |
| | | | | | | | Lough Foyle | | 20 | 50±18 | 4.7±1.6 | 0.07±0.02 | | | |
| CS/BBFraOT4 | EVHOE | Celtic Sea/Bay of Biscay | France | 1997 | 2016 | 4 | Thalassa II | Otter (GOV) | 20 | 30 ± 1 | 3.6±0.2 | 0.07±0.01 | NDB (92-07) DATRAS (08-15) | 10.7489/1958-1 | |
| BBIC(n)SpaOT4 | SP-North | Bay of Biscay and Iberian Coast | Spain | 1993 | 2014 | 4 | F deP Navarro | Otter (BACA) | 20 | 30 | | 0.05 | NDB | Not released, no DOI | |
| | | | | | | | Cornide Saavedra de | | 20 | 30 | 2.7±0.1 | | | | |
| | | | | | | | F deP Navarro | Otter (BACA) | 20 | 60 | 2.8±0.2 | | | | |
| BBIC(s)SpaOT1 | SP-ARSA | Bay of Biscay and Iberian Coast | Spain | 1990 | 2015 | 1 | F deP Navarro | Otter (BACA) | 20 | 60 | 5.6±0.2 | 0.1 ± 0.02 | NDB | Not released, no DOI | |
| | | | | | | | Cornide Saavedra de | | 20 | 60 | 5.6±0.4 | 0.1±0.01 | | | |
| BBIC(s)SpaOT4 | SP-ARSA | Bay of Biscay and Iberian Coast | Spain | 1997 | 2014 | 4 | F deP Navarro | Otter (BACA) | 20 | 60 | 5.5±0.3 | 0.09±0.02 | NDB | Not released, no DOI | |
| | | | | | | | Cornide Saavedra de | | 20 | 60 | 5.5±0.3 | 0.1±0.01 | | | |
| BBICPorOT4 | PT-IBTS | Bay of Biscay and Iberian Coast | Portugal | 2001 | 2014 | 4 | Capricornio, Noruega | Otter (NCT) | 20 | 29±3 | 3.1±0.4 | 0.05±0.01 | DATRAS | 10.7489/1963-1 | |
| WAScoOT3 | Rockall | Wider Atlantic | Scotland | 1999 | 2016 | 3 | Scotia III | Otter (GOV) | 20 | | | | DATRAS | 10.7489/1960-1 | |
| | | | | | | | | | | 30±3 | 3.4±0.4 | 0.07±0.01 | | | |
| WASpaOT3 | SP-PORC | Wider Atlantic | Spain | 2001 | 2015 | 3 | Vizconda de Eza | Otter (PBACA) | 20 | 24 ± 4 | 2.7±0.5 | 0.07±0.02 | NDB | Not released, no DOI | |
| GNSNetBT3 | BTS | Greater North Sea | The Netherlands | 1999 | 2016 | 3 | Isis, Tridens II | Beam (8m) | | | | | DATRAS | 10.7489/1967-1 | |
| | | | | | | | Isis | | 40 | 30±2 | 3.8±0.3 | 0.03 | | | |
| | | | | | | | Tridens II | | 40 | 34±11 | 4.5±1.4 | 0.04±0.01 | | | |

| | | | | | | | | | | | | | | |
|-----------|----------|-------------------|---------|------|------|---|---------------------|-----------|----|--------|---------|------|--------|----------------|
| GNSEngBT3 | BTS | Greater North Sea | England | 1990 | 2016 | 3 | Corystes | Beam (4m) | 40 | 29±3 | 3.7±0.6 | 0.01 | DATRAS | 10.7489/1966-1 |
| | | | | | | | Endevour | | 40 | 28±4 | 3.5±0.6 | 0.01 | | |
| GNSGerBT3 | BTS | Greater North Sea | Germany | 1998 | 2016 | 3 | Solea I | Beam (7m) | 40 | 30±3 | 3.5±0.5 | 0.03 | DATRAS | 10.7489/1965-1 |
| | | | | | | | Solea II | | 40 | 30 ± 2 | 3.3±0.3 | 0.02 | | |
| CSEngBT3 | BTS VIIa | Celtic Sea | England | 1993 | 2015 | 3 | Corystes, Endeavour | Beam (4m) | | | | | DATRAS | 10.7489/1964-1 |
| | | | | | | | Corystes | | 40 | 28 ± 5 | 3.8±0.5 | 0.02 | | |
| | | | | | | | Endevour | | 40 | 28 ± 4 | 3.5±0.6 | 0.01 | | |

584 *Table 1. List of individual surveys considered in the derivation of the Oslo/Paris convention (OSPAR) Groundfish Survey Monitoring and Assessment data*
585 *products. Survey acronyms reflect sub-region/country/gear/quarter, except CS/BB in the French EVHOE survey acronym to denote a survey that extends*
586 *across two sub-regions, the Celtic Seas and Bay of Biscay. Data product start and end years reflect the period when surveys were deemed sufficiently*
587 *established with consistent standardised methodology (Moriarty et al., 2017). NDB refers to national database. For this study we subset the data from 2004 –*
588 *2015 for continuous spatial coverage across the northeast Atlantic, the information on mesh size, haul duration, distance towed, and wing swept area reflect*
589 *the data included in this study from 2004-2015 where \bar{x} is the sample mean and s is the sample standard deviation.*

590

591 **Notes on fishing gear exceptions**

592 S = Standard Gear B = Bobbins used D = Double Sweeps I2 = Ground gear D with 16-inch bobbins R = Rockhopper

593 Grande Overture Vertical Trawl

- 594 1. Scotland uses R.V. Scotia III on five surveys WAScoOT3; CSScoOT4; CSScoOT1; GNSIntOT3; GNSIntOT1. For the west coast surveys (CSScoOT4 /
595 CSScoOT1/ WAScoOT3) they use a “S” and “I2” gear for to deal with rocky habitat. In North Sea surveys (GNSIntOT3; GNSIntOT1), Scotland uses
596 an “S” and a “B” exception.
- 597 2. Sweden uses a standard GOV (“S”) on R.V Argos and R.V. Mimer in both North Sea surveys (GNSIntOT3 and GNSIntOT1).
- 598 3. Denmark uses an “S” gear and an “R” exception in both surveys on R.V. Dana II in both North Sea surveys (GNSIntOT3 and GNSIntOT1).
- 599 4. England uses a standard GOV (“S”) gear in the North Sea (GNSIntOT3) on R.V. CEFAS Endeavour.
- 600 5. The Netherlands uses a standard GOV (“S”) gear in the North Sea (GNSIntOT1) on R.V. Tridens II. R.V. CEFAS Endeavour was used in quarter 1 by
601 Netherlands when Tridens broke down.
- 602 6. Norway uses an “S” gear and “D” exception on R.V. G.O. Sars and R.V. Johan Hjort in the North Sea (GNSIntOT3/ GNSIntOT1). When the R.V
603 Haakon Mosby has been used only a standard gear is noted.
- 604 7. France uses a GOV gear in the North Sea (GNSFraOT4) on R.V. Gwen Drez, no exception is noted, however, the gear is smaller than the standard
605 gear in the North Sea. France uses Thalassa II on two surveys CS/BBFraOT4 and GNSIntOT1. For the west coast surveys (CS/BBFraOT4) they use
606 ground gear “D” while in the North Sea surveys (GNSIntOT1), a standard gear is used.
- 607 8. Germany uses a standard gear on R.V Walther Herwig III in the North Sea (GNSIntOT3/ GNSIntOT1).
- 608 9. Ireland uses an “S” and “I2” gear for west coast survey (CSIREOT4) to deal with rocky habitat in line with Scotland on R.V. Celtic Explorer.

609 Beam Trawl

- 610 10. The Netherlands uses R.V. *Tridens II* and R.V. *Isis* in the GNSNetBT3 survey. Both ships use an 8m beam with a tickler but *Tridens II* has a different set
611 up to *Isis*.
- 612 11. Germany uses a 7m Beam trawl with a 5m tickler chain on R.V. *Solea II* during GNSGerBT3.
- 613 12. England uses a 4m Beam trawl during both her CSEngBT3 and GNSEngBT3 surveys on R.V. *Corystes* and CEFAS *Endeavour* in 2014 and 2015 with
614 the same rigging on both ships.
- 615 Rockhopper Trawl
- 616 13. The Rockhopper Otter Trawl is used by Northern Ireland in the CSNIrOT4 / CSNIrOT1 on R.V. *Corystes*.
- 617 Baka Trawl
- 618 14. Spain uses a Baka trawl on 3 surveys (BBIC(s)SpaOT4 / BBIC(s)SpaOT1 / BBIC(n)SpaOT4) on R.V. *Cornide de Saavedra*.
- 619 15. Spain uses a Porcupine Baka trawl on 1 survey (WASpaOT3) on R.V. *Vizconde de Eza*.
- 620 Norwegian Campelen Trawl
- 621 16. Portugal reports B and R gear exceptions on R.V. *Noruega*.

622 **Supplemental Material 1:**

623 **Table S1.1** Variance explained (%) and AIC scores for all models in all length classes. This
624 includes a folder containing the same information in Figure3 for every species/length
625 combination.

626 **Supplemental Material 2:** Can GAMMs differentiate effects of gear efficiency from spatial
627 and temporal variation in abundance in demersal fish?

628 **Supplemental Material 3:**

629 **File S3.1:** Example of R Scripts for fitting model

630 **File S3.2:** Example of R Scripts for generating data and fitting simulated models (S1
631 conditions only)