

Foreign Language Skills and Employment Status of European Natives: Evidence from Germany, Italy and Spain

Abstract

This article examines the relationship between foreign language skills and the employment status of natives in Germany, Italy and Spain. Using a probit model and data from Eurostat's Adult Education Survey 2011, this article studies the conditional correlations between knowledge of English and French as foreign languages, and the probability of being employed, comparatively, for men and women. The results reveal that skills in English increase the probability of being employed for men in the three countries, respectively, by 3.4, 4.3 and 5.2 per cent. Knowledge of English increases the probability of being employed for women in Germany and Italy – respectively, by 5.6 and 5.7 per cent – but not in Spain. The results also show that very good skills are associated with a higher probability of being employed than sufficient or good skills. The conditional correlation between knowledge of English and employment status for men is larger in countries where skills in this language are less common among the population, and where the unemployment rate is higher. Estimates for French are not statistically significant.

1. Introduction

This article addresses the relationship between foreign language skills and the working status of native Europeans, and it provides new, empirical results on this relationship. During the last decade, the official EU discourse about foreign language learning and teaching has gradually changed. While learning foreign languages was traditionally associated with an openness to other European cultures, nowadays the EU discourse on multilingualism emphasises the importance of foreign language skills for economic growth, competitiveness, mobility of labour, and employability. A working paper published in 2012 by the European Commission illustrates this trend:

Europe's vision for 2020 is to become a smart, sustainable and inclusive economy. Therefore, improving the outcomes of education and training and investing in skills in general—and language skills in particular—are important prerequisites to achieve the EU goal of increasing growth, creating jobs, promoting employability and increasing competitiveness. The ambition is to achieve better functioning of EU labour markets, to provide the right skills for the right jobs and to improve the

quality of work and working conditions. In this context, *foreign language proficiency is one of the main determinants of learning and professional mobility, as well as of domestic and international employability*. Poor language skills thus constitute a major obstacle to free movement of workers and to the international competitiveness of EU enterprises. [...] it is clear, however, that the benefits of improved language learning go well beyond the immediate economic advantages, encompassing a range of cultural, cognitive, social, civic, academic and security aspects (European Commission 2012: 4, italics added).

The EU, therefore, does not intend to neglect the cultural or cognitive aspects of language learning; quite simply the scope of EU language policy has been broadened. This change has gradually emerged since 2000 in different important policy documents such as the *Action Plan 2004-2006* (European Commission, 2003), the communication *A new strategic framework for multilingualism* (European Commission, 2005), and the communication *Multilingualism: An asset for Europe and a shared commitment* (European Commission, 2008). The change in the official discourse on multilingualism must be linked to two factors. The first one is strategic. Language policy, and in particular foreign language teaching, is viewed as an element contributing to the achievement of the general socio-economic objectives of the EU defined in the *Lisbon Agenda 2000-2010* (Krzyżanowski and Wodak, 2011), and to the achievement of the goals of the *Europe 2020 Agenda*.¹ The second factor is reactive. After the outbreak of the financial crisis in 2008 and the European sovereign debt crisis in 2011, the unemployment rate in the EU increased and reached a peak in 2013. It was 7 per cent in 2008 (6.6 per cent for men, 7.5 per cent for women), 10.9 per cent in 2013 (10.8 and 10.9 per cent for men and women, respectively), and then it decreased to 6.8 per cent in 2018 (6.6 per cent for men, 7.1 per cent for women). Large differences among countries exist. For example, in 2013, the unemployment rate was 5.2 per cent in Germany, 12.1 per cent in Italy, and 26.1 per cent in Spain. The youth unemployment rate is much higher than the average unemployment rate. The economic crisis severely hit the young. In 2008, the youth unemployment rate start-

¹ The Lisbon Agenda was a plan developed by the European Commission, aimed at making the EU “the most competitive and dynamic knowledge-based economy in the world, capable of sustainable economic growth with more and better jobs and greater social cohesion by 2010”. It was followed by Europe 2020, a ten-year strategy aiming at “smart, sustainable, inclusive growth” with greater coordination of national and European policy.

ed to grow quickly peaking in 23.8 per cent at the beginning of 2013, before receding to 15.2 per cent in 2018.

Against this backdrop, foreign language skills are viewed as a component of Europeans' human capital that can generate benefits in the domestic labour market, such as higher wages, or better employment opportunities. Reducing unemployment also matters for equity, because employment can be viewed as a dimension of social inclusion. Although the European Commission, as shown above, asserts that "foreign language proficiency is one of the main determinants of learning and professional mobility, as well as of domestic and international employability", few scientific papers address the question of the relationship between foreign language skills and employment, let alone employability. This article contributes to bridging this gap. To understand better why we focus on employment, a terminological note is needed. While the unemployment rate is clearly defined, the definition of what employability means is far from clear and many definitions coexist (McQuaid and Lindsay, 2005). The official definition used by the European Commission is the: "ability to gain initial employment, to maintain employment, and to be able to move around within the labour market", but it is not clear which indicators should be used to measure it. For this reason, the empirical studies available, as well as this article, concentrate on one dimension of employability only, that is, the employment status at a given moment.

The rest of this article is organised as follows: Section 2 discusses the relevant literature. Section 3 presents the database and our estimation strategy. Section 4 provides an overview of the linguistic skills of native-born citizens (henceforth "natives") in Germany, Italy and Spain, and it illustrates the characteristics of the sample. Section 5 presents the results of two econometric models. Section 6 provides a critical discussion of the results and an overall conclusion.

2. Literature Review

The majority of existing quantitative studies on the effects of language skills on labour market outcomes estimate earning differentials accruing to multilingual people, and for this reason they will not be discussed here (for recent overviews, see Gazzola *et al.*, 2016; Isphording, 2015; Zhang and Grenier, 2013; Chiswick and Miller, 2007).

Papers dealing with the effect of language skills on employment can be divided into two groups. The first group collects contributions that study the relationship between the employment opportunities of immigrants and the acquisition of good skills in the official language of the host country (Zorlu and Hartog, 2018; Yao and van Ours, 2015; Isphording and Otten, 2014; Chiswick and Miller, 2015; Maxwell, 2010; Aldashev *et al.*, 2009; Dustmann and Fabbri, 2003; Leslie and Lindley, 2001). Some authors examine the effects of acquiring proficiency in a regional, official, minority language, such as Catalan for people who move to Catalonia (Rendon, 2007), or English in South Africa (Cornwell and Inder, 2008).

The second group of papers focuses on foreign (or second) language skills and the employment status of the resident population, and is more relevant for the purposes of this article. The earliest study is by Vaillancourt (1988) who uses Canadian census data and a probit model to estimate the net impact of language skills on being employed or not in 1980 in Québec. The results showed that bilingualism in English and French increases the odds of being employed for women (men) by 6.5 per cent (1.8 per cent) for anglophones, 9.5 per cent (3 per cent) for francophones and 10.5 per cent (3.9 per cent) for allophones. Using an *ad hoc* data set in Switzerland, Grin *et al.* (2009, 2010) studied the relationship between employment and skills in English, French and German as second languages. They showed that, if the average wage increases by 5 per cent, the demand for monolingual workers decreases by 8.7 per cent, but the demand for multilingual workers decreases only by 3.7 per cent. Duncan and Mavisakalyan (2015) used the Data Initiative Survey to show that in some former Republics of the Soviet Union (i.e. Armenia, Azerbaijan and Georgia) skills in Russian language increase, *ceteris paribus*, the probability of employment by about 6 per cent for men and 9 per cent for women. Lindermann and Kogan (2013) employed data from the Estonian TIES survey and the Youth Transition Survey in Ukraine to study the role of language competency for labour market entry of the young in Estonia and Ukraine. The results showed that in Estonia knowledge of Estonian is important to the Russian-speaking minority in order to gain faster access to employment, whereas this is not the case for skills in Ukrainian in Ukraine.

There are four cross-national studies examining the relationship between employment and foreign language skills in the European Union. The first two studies used data from vacancy notices. Beadle *et al.* (2015) reported the results of 845 interviews with employers and employer organisations on the use and utility of foreign languages, the review of 3,632 online vacancy notices, and interviews with 522 employers. The authors show that a significant per-

centage of employers require an advanced level of foreign language skills. Fabo, Beblavy and Lenaerts (2017) investigated the economic importance of foreign language skills in the Visegrad group of countries (i.e. Czech Republic, Hungary, Poland, and Slovakia) using data obtained from key online vacancy boards and from the Wage Indicator survey, a continuous, voluntary, web-based survey of wages and working conditions. The results indicate that in the Visegrad region skills in English, and to a lesser extent German, are in high demand. The third study (Donado, 2017) investigated the conditional impact of knowing foreign languages on the likelihood of unemployment in different European countries using a repeated cross-section of more than 124,000 native residents (immigrants are excluded) aged 15 and over from 31 European countries. The data came from various waves of Eurostat's Eurobarometer surveys covering the period between 1990 and 2012. Using a linear probability model, the author showed that knowing a foreign language reduces the probability of being unemployed by at least 3.4 per cent, and this percentage is higher for women than men. This conditional impact is stronger for English (4.6 per cent), followed by German (2.1 per cent), Italian (2 per cent), and French (1.9), while the results for Spanish are not statistically significant. Substantial differences exist between the countries (Donado 2017: 275). Results obtained with an OLS regression in which linguistic distance is used as an instrumental variable (IV) show that the impact of foreign language skills on unemployment is significant only for English and German. Estimates of such an impact are higher than those obtained with the linear probability model. However, differences in the level of language skills were not examined.

The fourth study (Araújo *et al.* 2015) employed logistic regressions and data from Eurostat's Adult Education Survey (AES) 2011 to explore the relationship between knowledge of one or more foreign languages and the employment status of adult Europeans. The study reports a positive relationship between employment and knowledge of English in Germany, Greece, Italy, Latvia, Lithuania, Luxembourg, Malta, the Netherlands, Portugal and Slovenia. In Cyprus, Spain, and Finland proficiency in English (that is, very good language skills) is associated with a higher probability of being employed. People knowing at least some French are more likely to be employed in Malta, those who know German are more likely to be employed in Denmark, while Russian is associated with a higher probability of being employed in Bulgaria, Latvia, Lithuania and Poland. However, this paper shows three shortcomings. First, it reports the coefficients of the models used, but it does not estimate the related marginal effects. While this allows the authors to assess the sign of the existing relationships, no conclusion can be drawn about the magnitude of such relationships. Second, gender differ-

ences in the language-employment relationship are not explored. Finally, the relationship between employment and the level of language skills is incomplete, because it is examined only for English and only for a very good level of skills.

This article aims to deepen the research of Araújo *et al.* (2015).² Using probit models with detailed specifications, we provide estimates of the marginal effects of foreign language skills on the probability of being employed, for men and women separately, and for different levels of language skills. We focus on the domestic labour market of three EU countries – Germany, Italy and Spain – and on their most commonly spoken foreign languages, namely English and French. We select these countries for a number of reasons which are explained in Section 4. The available data do not offer sources of exogenous variation in language proficiency, which would allow us to identify those effects entirely unrelated to unobserved individual characteristics. Therefore, similar to other studies in language economics (e.g. Di Paolo and Tansel, 2019, and Yao and van Ours, 2019), this article focuses on existing conditional correlations between foreign languages and employment, rather than on the causal effect of the former on the latter. In both our models, however, we control for the main socio-economic determinants of employment in more detail than in Araújo *et al.* (2015), thus reducing the possibility of bias. Although we cannot draw conclusions of a causal nature, we consider that the estimates obtained are reliable from a qualitative point of view, and that they represent a useful contribution to future analyses.

3. Database and Estimation Strategy

This article uses data from the second edition of the AES, collected between 2011 and 2012 and published at the end of 2013. The survey covers the current 28 EU Member States, excluding Croatia, and some of the countries belonging to the European Free Trade Association (EFTA), such as Norway and Switzerland. The AES has a rich linguistic section recording data on the native language tongue and up to seven non-native language(s) of European residents aged 25-64. The data set covers 49 different languages. Furthermore, it provides data on respondents' level of proficiency in the two best-known non-native languages. More specifically, respondents are asked to self-assess their language skills using the following descriptors of competence:

² Our study is based on data from Eurostat Adult Education Survey (AES) 2011. The responsibility for all conclusions drawn from the data lies entirely with the authors.

- *Sufficient*: “I can understand and use the most common everyday expressions. I use the language in relation to familiar things and situations”.
- *Good*: “I can understand the essentials of clear language and produce simple texts. I can describe experiences and events and communicate fairly fluently”.
- *Proficient*: “I can understand a wide range of demanding texts and use the language flexibly. I master the language almost completely”.

A possible criticism concerns the reliability of subjectively self-assessed language skills. It is not possible to assess the accuracy of the AES data in this respect. However, research comparing self-assessment with teachers’ evaluations has shown that the former is more accurate if learners “respond to ‘can do’ statements that define concrete language use experiences that are familiar to the learners than if they are asked to use a proficiency scale with more abstract definitions of language skills” (Ross, 1998, quoted in Luoma, 2013: 4).

The survey also contains quite comprehensive information on the socio-economic status of the respondents, including their age, gender, family status (i.e. marital and parental status), level of education completed, and current labour status. Unfortunately, income data at the individual level are not available in the AES. Therefore, it is not possible to carry out an analysis of the relationship between languages skills and earning differentials.³

To study the effect of individual language proficiency on the employment status, we model the probability of being employed with a binary probit equation. In doing so, we interpret the explanatory variables as being able to influence an unobserved propensity for employment, for which we can only observe one binary output, i.e. employed or non-employed. By choosing the probit model instead of the logit model, as in Araújo *et al.* (2015), we are assuming this propensity to have standard normal distributed error terms.⁴ In the equation, we include survey variables that can influence the unobserved propensity for employment. Furthermore, taking advantage of the detailed data available, we specify the linguistic part of the equation

³ The income variable is defined as the decile of net monthly income of the household, including social benefits.

⁴ Probit models have already been used in the study of the relationship between language and employability, for instance in the case of Catalan in Catalonia (Rendon, 2007) and Québec (Vaillancourt 1988).

in two different ways. In the first specification, named Model 1, we treat language skills in a given language as a single dichotomous variable. In other words, we estimate the relationship between having at least some knowledge of a given foreign language and the probability of being employed, all other things being equal. In the second specification, Model 2, we introduce heterogeneity regarding language ability using the descriptors presented above. The first specification investigates whether knowledge of foreign languages, in general, is positively, conditionally correlated with the employment status of native individuals. The second specification deepens the analysis, exploring how specific levels of language proficiency relate to the employment status. Since foreign language skills are just one of the several variables that may impact on employment, we complete the analysis defining a set of socio-economic control variables, which remains unchanged in both models. Following the common practice in labour economics research, we control for age, experience, educational background, family status and regional effects.⁵ In the two models the conditional probability of employment is defined as follows:

Model 1:

$$\Pr[y_i = 1 | X_{ji}, FLA_i, FLB_i] = \Phi \left(\beta_0 + \sum_{j=1}^6 \beta_j X_{ji} + \gamma_1 FLA_i + \gamma_2 FLB_i + \varepsilon_i \right)$$

Model 2:

$$\begin{aligned} \Pr[y_i = 1 | X_{ji}, SFLA_i, SFLB_i] \\ = \Phi \left(\beta_0 + \sum_{j=1}^6 \beta_j X_{ji} + \sum_{l=1}^3 \theta_{1l} SFLA_i + \sum_{l=1}^3 \theta_{2l} SFLB_i + \varepsilon_i \right) \end{aligned}$$

In both equations, $y_i = \{0,1\}$ is a dummy variable for the employment status, equal to one if the individual is employed and to zero otherwise. Since we are using probit models, $\Phi(\cdot)$ is the cumulative distribution function for the standard normal distribution. In both models, we consider two different foreign languages, here labelled language *A* and language *B*, and we create specific linguistic variables for each of them. Each of these variables is constructed by merging the original data on the two best-known foreign languages reported by respondents.

⁵ See e.g. Aldashev, Gernandt and Thomsen (2009), Leslie and Lindley (2001), Rendon (2007) for similar specifications in comparable research questions.

Hence, the variable for language A (respectively, language B) records whether the respondent declares an ability to speak language A (language B) as a foreign language in general, no matter if it is his or her first or second best-known foreign language. In our analysis, proficiency in more than two foreign languages cannot be considered because the AES reports data on the level of skills just in the first and second foreign language. The specifications of Model 1 and Model 2 differ in how language skills are examined. In equation 1, FLA_i and FLB_i are two dummy variables representing the two foreign languages most frequently known by individual i in a given country, at any level of competence. Depending on one's language skills, an individual may not know any foreign language (both dummies equal to zero), only one language (FLA_i or FLB_i equal to one) or two languages (both FLA_i and FLB_i equal to one) at a sufficient, good or very good level. Model 2 does not aggregate the levels of competence in a single dummy for each language, but it analyses them separately. For this purpose, we define four different dummy variables for language A ($SFLA_i$) and language B ($SFLB_i$), and each dummy corresponds to a level of skill, including zero. Thus, in the equation 2, θ_{1l} (as well as θ_{2l}) captures the effect on employment of speaking language A (respectively, language B) at a skill level l , where $l = 1$ for sufficient, 2 for good and 3 for proficient. For each of the two languages, the reference level corresponds to respondents who declare no knowledge of the language considered (that is $l = 0$ for $SFLA_i$ and $SFLB_i$, respectively). As for the first model, also in the second equation, an individual may report knowing zero, one or two foreign languages.

The vector of variables $\sum_{j=1}^6 \beta_j X_{ji}$ is also common to the two models, and it controls for six dimensions, namely:

- *Age*: respondent' age, from 25 to 64.
- Age^2 : we control for non-linear age effects by including the square of *age*. Note that we cannot explicitly control for work experience, since this variable cannot be constructed from the data.⁶

⁶ Work experience is usually approximated by the age minus years of schooling minus six years of infancy. Note, however, that this measure would assume that all individuals work without breaks years after infancy and schooling. Since we want to estimate the relationship between individuals' language skills and their employment status, this approximation is likely to be harmful to our specification.

- *Married*: dummy variable indicating whether the respondent is married, including registered partnerships. Not being married includes widowed and not remarried, legally separated and not remarried, divorced, and single.
- *Child*: dummy variable indicating whether the respondent has at least one child aged less than 25 living at home.
- *Degree of urbanisation*: included with three alternative dummy variables indicating whether the respondent lives in a densely (*urb1*), intermediate (*urb2*) or thinly populated area (*urb3*). The reference category is the densely populated area (*urb1*). We control for the degree of urbanisation as a result of a lack of information on the geographical region in which a respondent lives. This is the only control available as a proxy for regional fixed effects.⁷
- *ISCED 1, ISCED2, ISCED3, ISCED4, and ISCED5*: dummy variables accounting for the highest level of education successfully completed by the respondent, according to ISCED classification.⁸ Education dummies cover primary/pre-primary (*ISCED1*), lower secondary (*ISCED2*), upper secondary (*ISCED3*), post-secondary non-tertiary (*ISCED4*) and tertiary-and-above (*ISCED5*) levels of education, respectively. The reference category for both models is a primary or lower level of education (*ISCED1*).⁹

4. Overview, Descriptive Statistics and Sample Characteristics

⁷ Information on the geographical regions would be a valuable source of variation in employment status across regions but it has been anonymised in the final version of the AES. See section 6 for a discussion.

⁸ The AES reports completed education in ISCED levels. ISCED stands for “International Standard Classification of Education”, a system developed by UNESCO to facilitate comparisons between the education systems of different countries. Omitting achieved education would introduce a relevant bias in our analysis, overestimating the impact of language skills on the explained variable.

⁹ The models specified by Araújo *et al.* (2015) control neither for work experience nor for the presence of children. Moreover, both variables capturing age and education effects are grouped into less specific macro-classes, and the degree of urbanisation is not taken into account.

We focus on native individuals in Germany, Italy and Spain for several reasons. First, these countries are among the most populated in Europe, and this is reflected in the size of the national statistical sample. Second, in the three countries the two most commonly spoken foreign languages are the same (e.g. English and French), and such languages are known by a significant percentage of the population. Third, they have only one official language at national level – local languages in Spain are co-official only in their respective regions – and none of their national or local languages coincides with English or French. Fourth, they give us the opportunity to investigate quite different contexts in terms of language knowledge and employment: in 2011, the employment rate was higher than the EU average in Germany, close to the average in Italy and below average in Spain.¹⁰ Finally, restricting the analysis to a few countries allows the study to be enriched with specific comments on their situations.

We focus on natives, which we define as the respondents who are both born in a given country and citizens of that country. This allows us to identify the actual natives correctly, as confirmed in Table 1 by the fact that the percentage of respondents declaring they speak German as a foreign language in Germany or Italian in Italy is close to zero. In Spain the fraction of natives who report Spanish as a foreign language mirrors the sociolinguistic situation of the country.¹¹ The article, therefore, deliberately excludes immigrants from the analysis, i.e. both EU-citizens abroad and people from third countries. First, immigrants' native language could correspond to a foreign language for natives (as English for immigrants from a former British colony), while at the same time immigrants may have limited competence in the official language of the host country. Second, the literature in language economics emphasises that immigrants can suffer from discrimination in the labour market of the host country for reasons that are independent of their skills in foreign languages. This can be the result of negative

¹⁰ Different factors may contribute to explaining these differences. Among others, let us mention the higher decentralisation in the wage setting in Germany, which considerably improved the competitiveness of the German economy in the last decade (Dustman *et al.*, 2014), as well as the importance of informal networks, such as family and friends, in the labour market in Italy and Spain (Naticchioni *et al.*, 2010, and Jaumotte, 2011).

¹¹ In some autonomous regions there may be an opposition between Spanish national identity on the one hand and regional identity on the other. Of the (few) Spaniards who claim to speak Spanish as a foreign language in the sample, six out of ten are native speakers of Catalan, followed by native speakers of Basque and Galician.

attitudes on the part of the native population towards their ethnic group, to their limited knowledge of the local dominant language (sometimes even simply to their accent) or a combination of the two factors (for a discussion see Chiswick and Miller, 2007, and Grin *et al.*, 2010). As a consequence, the analyses required for native and non-native individuals are conceptually different and, therefore, only the first group is considered here.

Table 1 shows the most commonly known foreign languages in the three countries studied. This table includes data from the first to the seventh foreign language known by respondents (if any).¹² English is more widely known in Germany than in Italy or Spain. Knowledge of French is more common in Italy than in Germany or Spain. Other languages are spoken by a lower (albeit not negligible) percentage of the population.

Insert Table 1 here

Table 2 reports respondents' self-reported levels of skill in English and French. Recall that variables on the level of skills are available only for the first and second foreign languages known by respondents and, therefore, the percentages in Table 1 and Table 2 can slightly differ. Data show that just a minority of individuals report being proficient in these two languages, a more significant part master them at a good level, while the majority reports knowing them at a sufficient level. Hence, contrary to what is commonly believed, fluency in foreign languages is not a basic skill of Europeans yet.

Insert Table 2 here

In building the outcome variable, and in a similar way to what was done by Araújo *et al.* (2015), we identify as employed ($y_i = 1$) individuals working either full-time or part-time. On the other hand, we consider as non-employed ($y_i = 0$) those who are not working, after

¹² Knowledge of more than two foreign languages is, however, uncommon: around 95% of the individuals in the three countries speak between zero and two foreign languages, with the two most spoken languages always being English and French. Adding the knowledge of a third language would increase this share from 95% to 99% of the population. The number of respondents for the third foreign language, however, is too small and does not allow for further investigation.

excluding students, permanently disabled individuals and respondents on compulsory military service¹³. To take gender differences into account we divide the sample between men and women.

While Table 1 and Table 2 refer to the whole population, Table 3 and Table 4 present the descriptive statistics of the sample, respectively, for men and women. Both tables report the share of respondents declaring a knowledge of English and French in general (this piece of information is used in Model 1), as well as the level of language skills in the two languages (these are examined in Model 2). Table 3 points out high levels of employment for men in all three countries. As expected, differences exist between countries: while Germany has the highest level of male employment (82 per cent), Spain shows the worst situation, with ten percentage points less (72 per cent), and Italy stands between them, but it is closer to Germany (78 per cent). The male sample maintains the characteristics of the whole population in terms of language knowledge and proficiency levels. English is widespread among German men (71 per cent), while fewer Italians (49 per cent) and fewer Spaniards (28 per cent) show some knowledge of it. French is better known in Italy than in the other two countries. Reported levels of proficiency are similar to those of the whole population for both languages.

Insert Table 3 here

Table 4 reports data for women. Differences in the percentage of employed people across countries are similar to those reported in Table 3 for men. One can note, however, that there is a substantial gender gap, as women are less likely to be employed than men. Such a gap is smaller in Germany and Spain (14 and 16 percentage points, respectively) than in Italy (21 percentage points). In the three countries, women's skills in English are similar to those of men, both in general and in terms of proficiency levels. Conversely, women tend to know

¹³ The non-employed category includes individuals who are (i) unemployed (ii) in retirement, early retirement or have given up business; (iii) fulfilling domestic tasks and (iv) other category of inactive person. This choice is to avoid removing conditions that are likely to depend on individual choices or gender. In particular, category (iii) is likely to be relevant when exploring the relationship between languages and employment for women: individuals doing domestic work represent almost 19 per cent of the female sample (and only 0.25 per cent of the male sample).

French better than men. Looking at the main socio-demographic variables, the male and female samples show similar proportions of married individuals in the countries examined. The percentage of women and men in the sample who have children is similar in Italy and Spain, while in Germany women are more likely to have children than men. We do not observe major gender differences in the degree of urbanization, while at country level Spain shows a higher concentration in thinly-populated areas. Finally, in Italy and Spain, women tend to be more educated than men, whereas the opposite is true in Germany, albeit to a lesser extent.

Insert Table 4 here

5. Estimates

From Table 5 to Table 8, we present the results of the estimates of Model 1 and Model 2. To assess both the sign and the magnitude of the conditional correlation between variables, we report in the table the marginal effects rather than the coefficients of the latent model. Marginal effects are computed at the mean (MEM). Hence, we evaluate the effect of foreign language skills on employment for the average man and woman in each country. Results for Model 1 and 2 are presented in Table 5 and Table 6 for the men's sample, while those for women are shown in Table 7 and Table 8. Recall that in all tables the reference levels for the degree of urbanization and education are, respectively, the densely populated area and ISCED 1. As to linguistic variables, the reference group in both models is composed of individuals who do not speak the language examined.

5.1 Results for men

Table 5 presents the results of model 1 for men. The vector of control variables shows the expected signs. Employment status is positively correlated with age and negatively correlated with its square. Being married and having children is positively correlated with the employment status of men in the three countries, which is consistent with the literature on the economic advantages linked to marriage (see Chun and Lee, 2001; Pollman-Schult, 2010). The degree of urbanization seems to have no effect, except for Spanish men living in thinly-populated areas, who seem to be more likely to be employed. As expected, education is positively and substantially rewarded in the labour market, with the probabilities of being employed generally increasing with higher levels of education.

Insert Table 5 here

Turning to the linguistic variables, results for Model 1 show that being able to speak English significantly increases the chances of employment for male individuals in all countries, albeit with differences between them. The marginal effects at the mean for French turn out to be statistically insignificant. Both findings are consistent with those of Araújo *et al.* (2015) for these countries. Although the authors do not compute the marginal effects, the coefficient of the variable denoting knowledge of English in their latent model is positive and statistically significant, whilst this is not the case for French. Focusing on the differences between the countries, we find that German men speaking English are 3.4 per cent more likely to be employed than their fellow countrymen who do not. This probability is equal to 4.3 per cent for Italian men and to 5.2 per cent higher in Spain. It is worth noting that the ability to speak English has a limited influence compared with that of the more general level of education. Having an upper secondary education (ISCED 3) enhances the probability of being employed by 31.9 per cent in Germany, 28 per cent in Spain and 22.9 per cent in Italy, that is, from five to more than nine times the magnitude of the respective effect of English skills in each country. At the same time, the conditional correlation between language knowledge and employment is not negligible for men, as it is comparable with the effect of an additional year in all countries.

The relationship between the different levels of language proficiency and employment status of men is presented in Table 6.

Insert Table 6 here

The marginal effects of non-linguistic control variables in Model 2 are almost identical to those in Model 1. As regards language skills, the marginal effects for French knowledge are statistically insignificant also when considering specific levels of proficiency. The results presented in Table 6 show that the absence of a conditional correlation between language skills in French and the employment status of men, in the three countries analysed, does not depend on the aggregation of skills' levels. The results for French, nevertheless, must be interpreted with caution because of the relatively low number of individuals in the sample who declare an ability to speak this language at a good, or very good, level.

As expected, a proficient level of skills in English is associated with a greater effect on employment status than sufficient or intermediate levels. There are, however, differences across

the countries in this respect. Good and very good skills in English increase the probability of being employed for German men by 3.8 and 5.4 per cent, respectively, while there seems to be no reward for a sufficient level. Italy and Spain show partially different results. In these two countries the conditional correlation between skills in English and respondents' employment status is statistically significant only for sufficient and proficient levels of skills, but not for the intermediate level. The probability of being employed for Italian men with sufficient skills in English is 4.3 per cent higher than for men who do not know this language, and it increases to 5.8 per cent if they are proficient in this language. For the Spanish men, the probability of being employed is slightly higher for men with very good language skills in English than for those with just sufficient skills (7 and 6.8 per cent, respectively).

5.2 Results for women

Table 7 reports the marginal effects for the variable of Model 1 for women. Again, the coefficients of the control variables have the expected sign. As in the case of men, age is positively correlated with employment status. The coefficient of the variable age squared is negative again. In contrast to the situation for men, and consistent with the existing literature in labour economics mentioned above, both marriage and children negatively affect the employment status of women. Interestingly, Italian women appear less likely to be non-employed when married, but not when having children living at home, while the opposite is true for German women. For Spanish women both marriage and having children are negatively correlated with their employment status. The degree of urbanization is correlated with employment, while education is still positively and largely rewarded in the labour market.

Insert Table 7 here

Knowledge of English has a positive and relevant effect on the employment status of German and Italian women. This effect is substantially higher than is the case for men: German and Italian women with English skills, respectively, are 5.6 and 5.7 per cent more likely to be employed than those not speaking it. Instead, knowing English has no significant effect on the employment status of Spanish women. As in the case of men, the conditional correlation between skills in French and employment status is not statistically significant in the three countries studied.

Table 8 presents the results of the regression of Model 2 for women. Also, in this case the estimates of MEM for non-linguistic control variables with Model 2 are very similar to those obtained when using Model 1.

Insert Table 8 here

Spanish women who are proficient in English have a 12 per cent higher probability of being employed than those not speaking it, while sufficient and good skills show no significant effect. This positive effect was “hidden” in Model 1 (in which the conditional correlation between a knowledge of English and employment status is not statistically significant), where differences in the level of linguistic skills are not considered. In the German and Italian labour market the magnitude of the conditional correlation between language skills and employment status varies with the level of proficiency. As in the case of men, for women higher competences are also associated with greater employment opportunities, but not all levels are rewarded. German women are 6.5 per cent more likely to be employed with a sufficient level of English, and 7.6 per cent more with a proficient level. For Italian women, the levels of English relevant for employment are sufficient and good, by 5.4 and 5.6 percentage points respectively. The coefficients for French are not statistically significant in the three countries.

6. Critical Discussion and Conclusions

This article examines the relationship between foreign language skills and the employment status of adult native citizens in three EU countries, namely, Germany, Italy and Spain. The pseudo- R^2 indicates that the models work better on the Italian data than the German and Spanish data. To assess whether the results obtained depend on the type of effects estimated, we have computed the average marginal effects for the same models presented in section 3. The results point in the same direction as the marginal effects at the mean. Similarly, we tested the robustness of our approach, based on the choice of a non-linear functional form, by estimating two linear probability models that include the same variables specified in the corresponding probit models. In the linear case too, the estimates show dynamics and magnitudes similar to those of the analysis presented in section 5.

We tried a specification of the probit models that used additional information in the field of education (e.g. humanities, social sciences, technology), but we found no improvements. A possible strategy to reduce heterogeneity is to use parents’ educational level as an instrumen-

tal variable instead of the observed level of foreign language knowledge. Nevertheless, this instrumental variable has often been criticised in the literature on language economics (see Chiswick and Miller, 2015: 242-243 for a detailed discussion), so the problem of disentangling the endogeneity of labour market outcomes and language proficiency has not yet been solved (Chiswick and Miller, 1995; Dustmann and van Soest, 2001). As a result, a relevant amount of variation in the employment status remains unexplained in our models. We expect that more variation in employment can be explained by differences in the types of job and/or sector of economic activity that are prevalent in the countries considered, e.g. export-oriented manufacturing industries, tourism or agriculture. In other words, part of the variation in the employment status could be explained by the different linguistic requirements in different types of occupation and/or sectors of economic activities. Tables A1 and A2 in the Appendix show the distribution of respondents' language skills in English and French by type of occupation, for men and women respectively. Tables A3 and A4 present the distribution of language skills in the most important economic sectors in terms of the percentage of individuals employed in the total target population. Data show that language skills tend to be more common in certain economic sectors and for certain types of occupation. However, the type of occupation and sector of economic activity are observable only for respondents who are employed and not for the non-employed, which are the two groups of interest to us. The influence of these two variables on employment status cannot be addressed in our models because they examine access to employment at a general level, rather than the specific occupations of jobholders. In order to evaluate the importance of this influence one should be able to control for the selection into labour market participation (this is especially relevant for women). Therefore, a thorough investigation of the relationship between occupation and language skills is a different research question that goes beyond the scope of this article and could be further explored in future analyses.

A methodological issue concerns endogeneity related to unobserved differences among regional labour markets. Foreign language skills may be better rewarded in regions that are closer to the border of neighbouring countries or that are more export-oriented. Indeed, in the three countries examined considerable regional differences exist as regards unemployment rates, for example between the ex-German Democratic Republic's (GDR) Länder and Bavaria or Baden-Württemberg regions, between the Italian Mezzogiorno and Lombardy, and between Andalusia and Catalonia in Spain. Further, it is possible that a very good knowledge of French has a significant, positive effect on the probability of being employed in border re-

gions, such as Saarland in Germany, or Piedmont in Italy. In this regard, we would expect that regional fixed effects could significantly improve the accuracy of the model. However, we cannot include precise regional fixed effects because no data on the respondents' region of residence is published in the AES. As already shown in Section 5, we have instead included the variables as proxies, but with little impact.

Notwithstanding these limitations, this article provides a new contribution to research on the economic value of language skills by presenting additional evidence on the positive relationship between foreign language skills and employment status. Using probit models and data from the Adult Education Survey 2011, we study the conditional correlations between knowledge of English and French as foreign languages and the probability of being employed in the three countries examined. The results reveal that skills in English increase the probability of being employed for men in Germany, Italy and Spain, respectively, by 3.4, 4.3 and 5.2 percent. For women in Germany and Italy, knowledge of English increases the probability of being employed by 5.6 and 5.7 per cent respectively, whereas in Spain the effect is not significantly different from zero. Increasing levels of English skills is associated with a higher probability of being employed. The results for French are not statistically significant in the three countries examined.

It is worth noting for men (and partially for women) the conditional correlation between knowledge of English and employment status is smaller in countries where skills in this language are more common and where the employment level is higher. This could be interpreted as a signal that skills in English tend to be less rewarded in the labour market as they become more widespread among the population, which is consistent with the very fundamental economic concept of scarcity. From this perspective, as English skills spread their value in the labour market will gradually fade (Grin 2001; Gazzola *et al.*, 2018). A second interpretation is that English language skills acquire greater value in labour markets with a lower employment rate because they allow individuals to have a further comparative advantage over competitors. The two interpretations, of course, are not mutually exclusive.

Finally, the results of this article can contribute to informing language policy decisions in education systems. As language skills are acquired mainly through education, establishing a positive relationship between this form of human capital and employment status stresses the importance of language education in Europe, and it provides evidence supporting the views of

different employers on this matter. For example, according to the study *Employers perceptions of graduate employability* carried out on 7,036 companies in Europe the majority of employers in the EU — except in Ireland, the UK and France — rates “very important” (33 per cent) or “rather important” (34 per cent) foreign languages skills when recruiting higher education graduates in their company (European Commission, 2010). The lack of sufficient foreign language skills on the labour supply side has been highlighted in various surveys or reports at national level, for example, in Austria (Tritscher-Archan, 2008), Denmark, France, Germany and Sweden (Bel Habib, 2011), Italy (Ministry of Labour, 2006), and the United Kingdom (Mulkerne and Graham, 2011).

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Tables

Table 1: Foreign languages known by natives aged 25-64 in Germany, Italy and Spain. Results in percentage

	Germany	Italy	Spain
Language			
English	72.0	46.2	31.2
French	19.8	23.2	13.0
German	0.6	4.9	2.0
Spanish	6.9	4.4	5.6
Italian	4.7	0.4	2.1
Russian	9.0	0.2	0.1

Individuals may report knowledge of more than one language up to seven languages. Results have been weighted so as to refer to the total population of each country.

Table 2: Level of skills in English and French of natives aged 25-64 in Germany, Italy and Spain. Results in percentage

	Germany	Italy	Spain
<i>English</i>			
No knowledge	28.3	54.1	69.5
Sufficient	36.2	29.7	12.1
Good	24.4	11.8	13.1
Proficient	11.1	4.4	5.3
<i>French</i>			
No knowledge	83.6	77.6	88.1
Sufficient	12.4	17.2	5.9
Good	2.9	3.8	4.4
Proficient	1.1	1.4	1.6

Only the first and the second best-known foreign languages are considered. Results have been weighted so as to refer to the total population of each country

Table 3. Descriptive statistics, Model 1 and 2. Men aged 25-64

Variable	Germany	Italy	Spain
employment	.82	.78	.72
age	46.72 (10.80)	46.36 (10.56)	44.69 (10.75)
married	.61	.67	.64
children	.33	.50	.48
<i>Degree of urbanization</i>			
densely populated	.48	.45	.43
intermediate	.36	.41	.25
thinly populated	.16	.14	.32
<i>Education</i>			
ISCED 1	.01	.06	.02
ISCED 2	.05	.33	.49
ISCED 3	.49	.39	.20
ISCED 4	.09	.05	0
ISCED 5	.36	.17	.29
<i>Language knowledge (general)</i>			
English	.71	.49	.28
French	.15	.20	.11
<i>Language knowledge (level)</i>			
<i>English</i>			
No knowledge	.29	.51	.72
Sufficient	.33	.30	.11
Good	.26	.14	.12
Proficient	.12	.05	.05
<i>French</i>			
No knowledge	.85	.80	.89
Sufficient	.12	.15	.05
Good	.02	.04	.04
Proficient	.01	.01	.02
Obs.	2,452	3,783	6,494

Results in share. Standard deviation, reported only for continuous variables, in brackets.

Table 4. Descriptive statistics, Model 1 and 2. Women aged 25-64

Variable	Germany	Italy	Spain
employment	.68	.57	.56
age	46.11 (10.49)	46.24 (10.54)	45.07 (10.70)
married	.63	.68	.67
children	.42	.51	.53
<i>Degree of urbanization</i>			
densely populated	.45	.46	.47
intermediate	.39	.41	.25
thinly populated	.16	.13	.28
<i>Education</i>			
ISCED 1	.02	.10	.02
ISCED 2	.08	.27	.43
ISCED 3	.50	.36	.22
ISCED 4	.08	.05	0
ISCED 5	.32	.22	.33
<i>Language knowledge (general)</i>			
English	.70	.49	.28
French	.19	.29	.13
<i>Language knowledge (level)</i>			
<i>English</i>			
No knowledge	.30	.51	.72
Sufficient	.36	.33	.12
Good	.24	.11	.12
Proficient	.10	.05	.04
<i>French</i>			
No knowledge	.81	.71	.87
Sufficient	.14	.22	.07
Good	.03	.05	.04
Proficient	.02	.02	.02
Obs.	2,713	3,995	7,010

Results in share. Standard deviation, reported only for continuous variables, in brackets.

Table 5. Probit regression. Men aged 25-64, Model 1

VARIABLES	(1) Germany	(2) Italy	(3) Spain
Age	0.049*** (0.006)	0.097*** (0.006)	0.049*** (0.005)
Age ² (× 10)	-0.007*** (0.001)	-0.012*** (0.001)	-0.007*** (0.001)
Married	0.096*** (0.017)	0.065*** (0.017)	0.101*** (0.016)
Child	0.033 (0.020)	0.037** (0.015)	0.062*** (0.015)
Intermediate area	-0.002 (0.0162)	0.003 (0.0143)	-0.008 (0.0163)
Thinly-populated area	0.013 (0.019)	-0.015 (0.022)	0.028* (0.015)
ISCED 2	0.217** (0.106)	0.109** (0.042)	0.163*** (0.062)
ISCED 3	0.319*** (0.097)	0.229*** (0.043)	0.280*** (0.063)
ISCED 4	0.342*** (0.100)	0.192*** (0.054)	
ISCED 5	0.387*** (0.098)	0.271*** (0.044)	0.373*** (0.062)
English (general)	0.034** (0.017)	0.043*** (0.015)	0.052*** (0.017)
French (general)	0.004 (0.022)	-0.023 (0.017)	-0.015 (0.021)
Observations	2,443	3,783	5,219
Pseudo-R ²	0.218	0.278	0.111

Marginal effects at the mean estimated for men aged between 25 and 64 years in Germany, Italy and Spain. Dependent variable: binary variable accounting for the individual employment status. Spanish data do not contain any observation with a level of education equal to ISCED 4. Standard errors in brackets *** p<0.01, ** p<0.05, * p<0.1.

Table 6. Probit regression. Men aged 25-64, Model 2

VARIABLES	(1) Germany	(2) Italy	(3) Spain
Age	0.048*** (0.006)	0.097*** (0.006)	0.049*** (0.005)
Age ² (× 10)	-0.006*** (0.001)	-0.012*** (0.001)	-0.006*** (0.001)
Married	0.098*** (0.017)	0.066*** (0.017)	0.101*** (0.016)
Child	0.031 (0.020)	0.036** (0.016)	0.062*** (0.015)
Intermediate area	-0.005 (0.016)	0.004 (0.014)	-0.007 (0.016)
Thinly-populated area	0.011 (0.019)	-0.015 (0.022)	0.027* (0.015)
ISCED 2	0.214** (0.106)	0.109** (0.042)	0.163*** (0.062)
ISCED 3	0.317*** (0.097)	0.229*** (0.043)	0.281*** (0.063)
ISCED 4	0.348*** (0.10)	0.193*** (0.054)	
ISCED 5	0.381*** (0.098)	0.270*** (0.044)	0.374*** (0.062)
sufficient English	0.030 (0.019)	0.043*** (0.016)	0.068*** (0.020)
good English	0.038* (0.023)	0.035 (0.024)	0.022 (0.023)
proficient English	0.054* (0.023)	0.059* (0.034)	0.070** (0.032)
sufficient French	0.002 (0.024)	-0.027 (0.020)	-0.032 (0.029)
good French	0.031 (0.045)	-0.024 (0.040)	-0.025 (0.037)
proficient French	0.029 (0.073)	0.023 (0.051)	0.047 (0.053)
Observations	2,433	3,783	5,215
Pseudo-R ²	0.223	0.278	0.112

Marginal effects at the mean estimated for men aged between 25 and 64 years in Germany, Italy and Spain. Dependent variable: binary variable accounting for the individual employment status. Spanish data do not contain any observation with a level of education equal to ISCED 4. Standard errors in brackets *** p<0.01, ** p<0.05, * p<0.1.

Table 7. Probit regression. Women aged 25-64, Model 1

VARIABLES	(1) Germany	(2) Italy	(3) Spain
Age	0.099*** (0.008)	0.114*** (0.008)	0.057*** (0.006)
Age ² (× 10)	-0.012*** (0.001)	-0.014*** (0.001)	-0.007*** (0.001)
Married	-0.032 (0.021)	-0.149*** (0.021)	-0.092*** (0.016)
Child	-0.095*** (0.023)	0.004 (0.021)	-0.030* (0.016)
Intermediate area	0.009 (0.020)	-0.001 (0.019)	-0.008 (0.017)
Thinly-populated area	0.001 (0.027)	-0.006 (0.028)	-0.0158 (0.017)
ISCED 2	0.183** (0.084)	0.162*** (0.034)	0.209*** (0.039)
ISCED 3	0.313*** (0.078)	0.371*** (0.035)	0.410*** (0.041)
ISCED 4	0.374*** (0.083)	0.339*** (0.049)	
ISCED 5	0.465*** (0.078)	0.502*** (0.036)	0.557*** (0.040)
English (general)	0.056** (0.022)	0.057*** (0.020)	0.011 (0.018)
French (general)	-0.036 (0.0260)	-0.026 (0.020)	0.003 (0.022)
Observations	2,699	3,995	5,685
Pseudo-R ²	0.108	0.200	0.136

Marginal effects at the mean estimated for women aged between 25 and 64 years in Germany, Italy and Spain. Dependent variable: Binary variable accounting for the individual employment status. Spanish data do not contain any observation with a level of education equal to ISCED 4. Standard errors in brackets *** p<0.01, ** p<0.05, * p<0.1.

Table 8. Probit regression. Women aged 25-64, Model 2

VARIABLES	(1) Germany	(2) Italy	(3) Spain
Age	0.099*** (0.008)	0.114*** (0.008)	0.057*** (0.006)
Age ² (× 10)	-0.012*** (0.001)	-0.014*** (0.001)	-0.007*** (0.001)
Married	-0.030 (0.022)	-0.149*** (0.021)	-0.093*** (0.016)
Child	-0.096*** (0.023)	0.004 (0.021)	-0.028* (0.016)
Intermediate area	0.010 (0.020)	-0.001 (0.019)	-0.008 (0.018)
Thinly-populated area	0.001 (0.027)	-0.005 (0.028)	-0.014 (0.017)
ISCED 2	0.177** (0.085)	0.164*** (0.034)	0.210*** (0.039)
ISCED 3	0.309*** (0.079)	0.372*** (0.035)	0.411*** (0.041)
ISCED 4	0.372*** (0.085)	0.340*** (0.049)	
ISCED 5	0.458*** (0.080)	0.500*** (0.036)	0.554*** (0.040)
sufficient English	0.065*** (0.024)	0.054** (0.021)	-0.018 (0.023)
good English	0.036 (0.029)	0.056* (0.032)	0.014 (0.025)
proficient English	0.076** (0.038)	0.076 (0.048)	0.121*** (0.037)
sufficient French	-0.029 (0.031)	-0.035 (0.022)	0.030 (0.029)
good French	0.008 (0.053)	0.026 (0.042)	-0.0042 (0.035)
proficient French	-0.080 (0.083)	-0.042 (0.062)	-0.067 (0.057)
Observations	2,689	3,995	5,680
Pseudo-R ²	0.107	0.200	0.137

Marginal effects at the mean estimated for women aged between 25 and 64 years in Germany, Italy and Spain. Dependent variable: binary variable accounting for the individual employ-

ment status. Spanish data do not contain any observation with a level of education equal to ISCED 4. Standard errors in brackets *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Appendix

Table A1. Distribution of skills in English and French (at any level), by occupation type. Men aged 25-64. Results in percentage

Occupation Type	Percentage of employed men	Germany		Italy		Spain	
		English	French	English	French	English	French
0. Armed forces (professionals)	1.5	100.0	15.0	72.0	10.7	37.3	19.6
1. Managers, senior officials, legislators	6.4	88.1	17.9	60.0	21.7	38.2	18.7
2. Professionals	16.5	95.3	33.3	77.3	26.1	59.6	22.7
3. Technicians and associate professionals	18.6	83.6	16.9	69.8	24.5	48.5	16.3
4. Clerical support workers	7.1	79.7	20.9	61.7	25.7	36.0	13.7
5. Service and sales workers	10.4	69.3	10.1	47.6	16.2	31.4	7.8
6. Skilled agricultural, forestry and fishery workers	3.8	57.1	7.1	31.1	16.7	6.9	6.0
7. Craft and related trades workers	20.8	57.6	4.8	31.5	14.5	17.4	6.1
8. Plant and machine operators and assemblers	9.2	55.4	6.6	35.1	10.6	14.8	5.9
9. Elementary occupations	5.7	41.5	7.3	22.5	8.8	15.0	5.2
Total	100.0						

Occupation types are reported in ISCO-08 classification at 1-digit level.

Table A2. Distribution of skills in English and French (at any level), by occupation type. Women aged 25-64. Results in percentage

Occupation Type	Percentage of employed women	Germany		Italy		Spain	
		English	French	English	French	English	French
0. Armed forces (professionals)	0.1	100.0	0.0	75.0	50.0	0.0	0.0
1. Managers, senior officials, legislators	3.4	93.2	36.4	68.1	46.4	40.5	18.4
2. Professionals	22.8	91.8	45.0	73.4	31.3	55.6	21.9
3. Technicians and associate professionals	22.2	77.9	19.8	72.0	32.3	46.1	20.0
4. Clerical support workers	14.8	85.7	20.1	71.4	38.9	47.0	16.9
5. Service and sales workers	20.5	61.4	7.9	45.4	25.7	23.0	10.8
6. Skilled agricultural, forestry and fishery workers	1.3	40.0	0.0	37.1	25.7	7.8	3.1
7. Craft and related trades workers	4.0	62.5	12.5	33.0	14.8	15.6	6.5
8. Plant and machine operators and assemblers	1.5	61.5	0.0	24.5	20.8	15.8	5.3
9. Elementary occupations	9.4	43.1	5.0	28.0	23.0	8.8	3.9
Total	100.0						

Occupation types are reported in ISCO-08 classification at 1-digit level

Table A3. Distribution of skills in English and French (at any level) in the five most important sectors of economic activity. Men aged 25-64. Results in percentage

Sectors	Percentage of employed men	Germany		Italy		Spain	
		English	French	English	French	English	French
C. Manufacturing	20.7	75.4	12.7	49.7	15.8	28.9	9.2
F. Construction	11.5	53.3	5.7	34.3	13.0	15.1	6.0
G. Wholesale and retail trade	12.8	73.3	7.8	47.0	19.3	29.1	9.2
H. Transportation and storage	7.1	66.7	12.5	52.0	16.3	24.5	9.5
O. Public administration and defence	9.1	83.8	13.8	60.5	19.9	34.8	15.8
Total	61.2						

Sectors of economic activities are reported in NACE Rev.2 classification at 1-digit level.

Table A4. Panel B. Distribution of skills in English and French (at any level) in the five most important sectors of economic activity. Women aged 25-64. Results in percentage

Sector	percentage of employed women	Germany		Italy		Spain	
		English	French	English	French	English	French
C. Manufacturing	9.6	69.6	15.2	55.5	27.0	28.0	10.9
G. Wholesale and retail trade	15.6	66.7	11.1	54.8	31.6	34.5	10.3
M. Professional, scientific and technical activities	6.3	91.7	36.9	76.1	32.6	56.7	18.0
O. Public administration and defence	8.0	83.2	28.0	64.4	41.1	37.2	22.2
Q. Human health and social work activities	17.2	75.0	19.6	63.8	29.5	34.6	19.3
Total	56.7						

Sectors of economic activities are reported in NACE Rev.2 classification at 1-digit level.