



The long and short of it: a comprehensive assessment of axial length estimation in myopic eyes from ocular and demographic variables.

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1 Title: The long and short of it: a comprehensive assessment of axial length estimation in myopic eyes
2 from ocular and demographic variables.

3 Running title: Estimating axial length for myopia management

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27 **ABSTRACT**

28 Background/Objectives

29 Axial length, a key measurement in myopia management, is not accessible in many settings. We
30 aimed to develop and assess machine learning models to estimate the axial length of young myopic
31 eyes.

32 Subjects/Methods

33 Linear regression, symbolic regression, gradient boosting and multilayer perceptron models were
34 developed using age, sex, cycloplegic spherical equivalent refraction (SER) and corneal curvature.

35 Training data were from 8135 (28% myopic) children and adolescents from Ireland, Northern Ireland
36 and China. Model performance was tested on an additional 300 myopic individuals using traditional
37 metrics alongside the estimated axial length vs age relationship. Linear regression and receiver
38 operator characteristics (ROC) curves were used for statistical analysis. The contribution of the
39 effective crystalline lens power to error in axial length estimation was calculated to define the
40 latter's physiological limits.

41 Results

42 Axial length estimation models were applicable across all testing regions ($p \geq 0.96$ for training by
43 testing region interaction). The linear regression model performed best based on agreement metrics
44 (mean absolute error [MAE]=0.31mm, coefficient of repeatability=0.79mm) and a smooth,
45 monotonic estimated axial length vs age relationship. This model was better at identifying high-risk
46 eyes (axial length >98th centile) than SER alone (area under the curve 0.89 vs 0.79, respectively).
47 Without knowing lens power, the calculated limits of axial length estimation were 0.30mm for MAE
48 and 0.75mm for coefficient of repeatability.

49 Conclusions

50 In myopic eyes, we demonstrated superior axial length estimation with a linear regression model
51 utilising age, sex and refractive metrics and showed its clinical utility as a risk stratification tool.

52 **INTRODUCTION:**

53 Myopia is typically caused by abnormal axial eye growth between childhood and young adulthood¹
54 and this excessive ocular elongation is strongly linked to irreversible vision loss among myopic
55 adults.² This is one of the reasons why axial length has been proposed as the key metric for assessing
56 myopia control treatment efficacy.³

57 To maximise accessibility for patients and efficiently utilise expertise and workforce, the
58 management and treatment of myopia progression will likely take place in primary and community
59 eye care settings.⁴ However, unlike specialised and secondary eye care settings, access to biometric
60 devices to measure axial length in primary and community care settings is limited. Recent guidelines
61 and research publications have emphasised the value of axial length in managing myopia
62 progression.⁵⁻⁷ Thus, the inability to measure axial length is a potential barrier to the uptake of
63 myopia control treatments and limits a clinician's ability to identify patients most in need of
64 treatment.

65 Spherical equivalent refractive error (SER) and axial length are intrinsically linked, with SER related to
66 the mismatch between the axial length and focal power of the eye, the latter being predominantly
67 derived from the cornea and crystalline lens. Previous research has shown that axial length can be
68 estimated from SER alone or a combination of SER and corneal curvature.⁸⁻¹² However, challenges in
69 accurately estimating axial length from SER remain,^{3,13,14} with published models⁸⁻¹¹ generally
70 performing below a clinically desirable level and failing to account for the impact of age and sex on
71 axial length,¹¹ potentially non-linear associations^{13,15} and alterations in the axial length to SER
72 relationship that occur with age.^{11,16}

73 Machine learning aims to develop predictive models from data and spans a range of techniques from
74 linear regression to more complex deep learning models.¹⁷ The full potential for machine learning
75 techniques to improve axial length estimation models has not been explored. We aimed to: 1)
76 develop and compare new machine learning models using age, sex, SER and corneal power to

77 estimate axial length, 2) compare these models to published algorithms and evaluate their
78 application in the clinical management of myopia and 3) explore the physiological limits of axial
79 length estimation, using the known biological sources of error in the estimation of axial length from
80 refractive variables.

81 **Subjects and Methods**

82 *Model development*

83 Four machine learning approaches were used to develop new models for estimating axial length:
84 linear regression, symbolic regression, gradient boosting and multilayer perceptron. Multilayer
85 perceptron is a form of artificial neural network, while gradient boosting is an ensemble method that
86 combines a series of weak machine learning models. Neither model produces a defined equation.
87 Symbolic regression iteratively searches mathematical relationships to determine the simplest
88 equation for accurately describing relationships within a dataset. For each model, input variables
89 were cycloplegic SER (dioptres), average corneal radius of curvature (mm), biological sex, age (years)
90 and, due to the known logarithmic relationship between axial length and age,¹⁸ the base 10
91 logarithm of age. Corneal curvature was included as widely available autorefractors measure both
92 SER and corneal curvature.

93 The standard multiple linear regression model was built using R (R Foundation for Statistical
94 Computing, Austria) and interaction terms were tested and retained if they improved prediction
95 performance on validation data (see below). Symbolic regression, gradient boosting and multilayer
96 perceptron models were implemented using the Scikit-learn Python library.¹⁹ For the gradient
97 boosting model, the number of estimator parameters was fixed at 600 and learning rate, maximum
98 depth and sub-sampling fraction parameters were fixed at 0.02, 2, and 0.2, respectively. For the
99 multilayer perceptron model, the maximum number of iterations was fixed at 1000 and the logistic
100 function used as the activation function for hidden layers. Symbolic regression was implemented
101 using the GP-GOMEA library²⁰ with the number of generations fixed at 10.

102 The four newly developed machine learning models were compared to published algorithms, which
103 used either the Gullstrand schematic eye¹⁰ or least-squared techniques such as simple²¹ or multiple
104 linear regression.^{8,9,11,12} No previous algorithms used sex, while only Queirós et al. included age.¹²

105 *Training data*

106 Training data were derived from three epidemiological studies, the Ireland Eye Study,²² the Northern
107 Ireland Childhood Errors of Refraction (NICER) study^{23,24} and a refractive error study from the
108 Zhejiang province of China.²⁵ The Ireland Eye Study (IES) included a population-based sample of 1626
109 children aged 6-7 or 12-13 years. The NICER study is a population-based cohort study of children in
110 Northern Ireland with data from 2 424 examinations of 1 068 children aged 6-20 years. The Zhejiang
111 refractive error survey was a school-based cross-sectional study of 3 429 participants aged 5-18
112 years. The training data additionally included data from myopic children aged 6-19 years enrolled in
113 studies, including myopia control clinical trials, at the Centre for Eye Research Ireland (CERI; 806
114 visits from 240 participants).²⁶ Participant region was defined as either the island of Ireland (referred
115 to as Ireland) or China. Across all studies, informed consent was obtained from legal guardians, SER
116 was measured using cycloplegic autorefraction and axial length, corneal curvature and anterior
117 chamber depth (ACD) measured using an ocular biometer (details published elsewhere²²⁻²⁷). For
118 model training, the training data were randomly divided, with 95% used for model training and the
119 remaining 5% for validation, where validation data were used for iterative assessment of model
120 performance during training to avoid overfitting.

121 *Testing data*

122 Final models were tested on myopic participants (SER \leq -0.50 dioptres) only, using a separate,
123 randomly selected subset of 150 participants of the cross-sectional Zhejiang study and 150
124 participants of the CERI clinical trial, the Myopia Outcome Study of Atropine in Children (MOSAIC;
125 ISRCTN: 36732601). The myopic testing population reflects the proposed clinical application of an
126 axial length estimation tool, which is to assess and monitor the axial length of myopic children and

127 adolescents. MOSAIC participants were enrolled in a randomised placebo-controlled clinical trial of
128 nightly atropine 0.01% eye drops and data were available at baseline and the 12-month follow-up.
129 At the 12-month follow-up, 2/3 of MOSAIC participants were using atropine 0.01% eye drops, which
130 could reflect a clinical setting.

131 *Applicability of models*

132 To investigate whether models trained on European data could be applied to Asian populations and
133 vice versa, models were trained using either all data combined, or data from Ireland or China alone.
134 The performance of these models was then assessed, separately, using testing data from either
135 Ireland, China or both.

136 *Model performance*

137 We used a mixture of traditional and non-traditional statistics for assessing model performance:

- 138 1. Agreement between estimated and actual axial length, including 12-month change.
- 139 2. Ability to identify high-risk eyes with long axial length or fast axial length growth
- 140 3. Monotonicity and smoothness of axial length estimations as a function of age.

141 *Agreement between estimated and actual axial length*

142 Agreement between estimated and actual axial length was assessed by analysing the mean bias and
143 standard deviation of the prediction errors, the mean absolute error (MAE), the concordance
144 correlation coefficient,²⁸ which measures the correlation along a line of perfect agreement, and the
145 coefficient of repeatability. The coefficient of repeatability indicates the interval around the
146 estimated axial length in which the true measurement would lie with 95% confidence and was
147 calculated as $2.77 \times$ within-subject standard deviation.²⁹ Change in axial length was estimated by
148 subtracting the estimated axial length at baseline from that at the 12-month follow-up.

149 *Identification of high-risk eyes*

150 High-risk eyes were defined as eyes with axial length in the top 2% for the participant's age and sex.
151 To calculate this threshold, the 98th centiles of axial length were extracted from two reports of
152 European participants^{18,30} and a weighted cubic spline fitted to these data to generate age- and sex-
153 specific 98th centiles (Supplementary Figure 1). Fast axial length progression was defined as
154 >0.20mm/year based on the mean axial length growth among myopic eyes in the NICER study.³¹

155 *Monotonicity and smoothness*

156 An estimated axial length that, in the absence of corneal or SER changes, decreases or varies
157 considerably with increasing age is not clinically plausible. We, therefore, examined the
158 monotonicity and smoothness of axial length estimations as a function of age by creating curves in
159 which axial length was estimated for each year of age between 6 and 20 years, while all other
160 variables were held constant. The first derivative (rate of change) of the age curve was calculated,
161 and monotonicity was assessed by testing whether the first derivatives were all ≥ 0 (a negative value
162 would indicate decreasing axial length with age) and smoothness was evaluated by comparing
163 deviations of the first derivate from a fitted smooth spline curve.

164 *Model extension*

165 Anterior chamber depth (ACD) alters effective lens power and should, therefore, account for some
166 of the variance in axial length not explained by SER or corneal power. We investigated whether
167 adding ACD improved axial length estimation. ACD was either estimated from known relationships
168 with age,³² to determine whether explicitly defining this relationship in the model could improve its
169 performance, or actual ACD measured from biometry was used.

170 *Impact of sex, age and spherical equivalent on axial length estimation*

171 To illustrate the benefits of including sex in the algorithm, we trained additional models using either
172 female or male training data and tested these models on either males, females or both sexes

173 combined in the myopic testing population. The influence of age and SER on axial length estimation
174 was assessed by quantile regression using axial length estimation error as the outcome.

175 *Limits of axial length estimation*

176 Errors in axial length estimation largely arise from the unmeasured effective crystalline lens power.
177 When the true axial length, corneal power and SER are known, the refractive power of the internal
178 ocular structures (internal refractive power; IRP) can be calculated and, by comparing against age-
179 and sex-matched emmetropic eyes, the contribution of the IRP to SER calculated.³³ Assuming IRP
180 and corneal power are independent, the variance of the IRP contribution to SER can be used to
181 explore the expected variance in axial length estimation error arising from IRP alone.

182 *Spectacle refraction*

183 As cycloplegic autorefraction is not always feasible, we extended the axial length estimating model
184 to work with presenting spectacle refraction by introducing an intermediate step that uses multiple
185 linear regression to estimate cycloplegic autorefraction from spectacle refraction and age. Only data
186 from CERI clinical trials contained both presenting spectacle refraction and cycloplegic
187 autorefraction. Undercorrected participants were excluded by removing examinations with
188 presenting corrected visual acuity >0.1 logMAR.

189 *Statistical analysis*

190 Only right eye data were used, tests are two-sided and significance was set at $p < 0.05$. Model
191 performance metrics were compared using linear regression, adjusting for the model type. The
192 effects of training region, testing region and including ACD on model performance were assessed by
193 adding these as covariates. A linear mixed model was used to estimate cycloplegic autorefraction
194 SER from spectacle refraction using random intercept terms to account for any within-subject
195 correlations. Performance metrics were calculated within Python (v3.10.4, Python Software

196 Foundation, USA), and linear and quantile regression was performed with R (v4.3.0, R Foundation for
197 Statistical Computing, Austria).

198 **RESULTS**

199 *Population characteristics*

200 Table 1 provides an overview of participants characteristics in the testing and training data sets. In
201 the training data, participants from China were younger on average, but otherwise had similar
202 characteristics to participants from Ireland. As expected, participants in the testing data were more
203 myopic than the training data.

204 *Applicability of models*

205 Table 2 shows the performance of the four machine learning models when trained and then tested
206 on the combined data or from Ireland or China, separately. While there were significant main effects
207 of model training and testing regions on linear regression, the effect of training region on all model
208 metrics did not vary significantly across testing region and vice versa (interaction for all $p \geq 0.97$).

209 *Model performance*

210 We compared the performance of these newly developed axial length estimation models against
211 previously published algorithms, when applied to the myopic testing data, at either a single visit
212 (cross-sectional, Figure 1a) or when estimating the 12-month change in axial length (longitudinal
213 [Irish data only], Figure 1b). Due to the number of algorithms, data are presented as scatter plots of
214 agreement metrics from Table 2 (all regions), with each plot including a measure of agreement
215 (concordance correlation coefficient, mean bias) and a measure of the spread of estimation errors
216 (coefficient of repeatability, standard deviation of errors), which are often related, on each axis. A
217 good axial length estimator would maximise exact agreement while minimising the spread of errors.
218 Compared to cross-sectional performance, the concordance correlation coefficient for estimation of
219 axial length change was poorer across all models, possibly because rounding of SER to the nearest

220 0.125 dioptre results in a proportionally larger loss of information. The newly developed models
221 outperformed all previously published models, with the linear regression model performing best or
222 equal best across both cross-sectional and longitudinal axial length estimation.

223 *Smoothness and monotonicity*

224 We next assessed the modelled relationship between estimated axial length and age (Figure 2). The
225 gradient boosting model was monotonic, but not smooth, explaining the poor longitudinal
226 performance (Figure 1b). The multilayer perceptron model was smooth, but not monotonic,
227 predicting that highly myopic eyes of young adult females shorten over time. The linear and
228 symbolic regression models fulfilled the pre-specified criteria of monotonicity and smoothness,
229 indicating the modelled axial length vs age relationship was clinically plausible.

230 Figure 3 shows a clinically relevant assessment of the top performing linear regression model
231 (Equation 1). Prediction errors were approximately normally distributed around a small mean bias.
232 Despite a reasonable level of uncertainty surrounding estimated axial length, 86% of estimations
233 were within $\pm 0.50\text{mm}$ of the actual axial length for cross-sectional data and 82% were within
234 $\pm 0.10\text{mm}$ on longitudinal testing. The linear regression model was better than SER alone at
235 identifying high-risk eyes with axial length in the top 2% for age and sex (area under the curve [AUC]
236 0.89 vs 0.79, respectively), correctly identifying 99/121 (82%) high-risk eyes compared to only
237 88/121 (73%) using the optimum SER threshold of -3.37 D. However, the linear regression model
238 was not better than SER at identifying eyes with fast axial length progression $>0.2\text{mm/year}$ (Figure
239 3).

$$240 \quad AL = 5.472 - 0.069 \times age + 3.060 \times \log_{10}(age) - 0.265 \times sex -$$
$$241 \quad 0.135 \times SER + 2.019 \times CR - 0.072 \times SER \times CR$$

242 **Equation 1** Top-performing linear regression equation where AL is the axial length in mm, age is in
243 years, sex=0 for males and 1 for females, SER is spherical equivalent refractive error in dioptres, and
244 CR is the corneal radius of curvature in mm.

245 *Model extension with anterior chamber depth*

246 We investigated whether model performance could be improved by including either an estimated
247 ACD, derived from age, or actual ACD. The performance metrics of the original and ACD-extended
248 models are shown in Supplementary Table 1 and, for the linear regression model, in Figure 3.
249 Compared to models without ACD, adding actual ACD significantly improved axial length estimation
250 (MAE difference=-0.03mm, p=0.001; coefficient of repeatability difference=-0.08mm, p<0.001),
251 whereas adding estimated ACD did not (MAE difference=-0.001, p=0.99; coefficient of repeatability
252 difference=0.003mm, p=0.96).

253 *Model performance by sex, age and spherical equivalent*

254 Compared to the model in Equation 1, the performance of linear regression models were worse
255 when trained using either only males or females and then tested on participants of the opposite sex
256 (Supplementary Table 2). Models trained on male data overestimated female axial length (mean
257 bias=-0.19mm) and models trained on female data underestimated male axial length (mean
258 bias=0.33mm).

259 We used quantile regression to assess whether the linear regression model performance varied with
260 age or SER (Supplementary Table 3, Supplementary Figure 2). Quantile slopes were positive at the
261 2.5% and 50% centile for SER and the 2.5% centile for age, indicating median bias tended to become
262 more positive with decreasing myopia and that axial length estimation was less precise for highly
263 myopic (<-8 D) and older participants.

264 *Limits to axial length estimation*

265 In the myopic testing population, the distribution of the IRP contribution to SER was Gaussian with a
266 standard deviation of 0.91 dioptres. From this distribution (mean shifted to 0), we randomly
267 sampled 10,000 observations (referred to as IRP error) and calculated the estimated axial length
268 with and without IRP error added to the SER parameter, for ages 6-20 years, males and females,
269 separately, and corneal curvature held constant at 7.76mm. The difference between the two
270 estimations represents the axial length estimation error attributable to the variance in IRP around
271 the population mean. From this, the minimum obtainable coefficient of repeatability was ± 0.75 mm,
272 and MAE was 0.30mm, comparable to the linear regression model's ± 0.80 mm and 0.31mm,
273 respectively.

274 *Spectacle refraction*

275 Of the 806 CERI clinical trial training data, 731 (90.7%) examinations had corrected visual acuity of
276 0.10 logMAR or better and were used to model cycloplegic autorefraction from spectacle refraction
277 with a linear mixed model (Equation 2).

$$278 \text{SER}_{\text{cyclo autoref}} = -0.153 + 0.955 \times \text{SER}_{\text{spectacle}} - 0.028 \times \text{age} - 0.013 \times \text{SER}_{\text{spectacle}} \times \text{age}$$

279 **Equation 2** Formula for estimating cycloplegic autorefraction (cyclo autoref) spherical equivalent
280 from spectacle spherical equivalent (D) and age (years).

281 Compared to spectacle refraction alone, the formula in Equation 2 improved the estimation of
282 cycloplegic autorefraction SER in the training dataset (MAE improved from 0.61 D to 0.45 D;
283 concordance correlation coefficient improved from 0.931 to 0.945). This allows an estimate of axial
284 length in the absence of cycloplegic autorefraction, where the measured cycloplegic autorefraction
285 SER in Equation 1 is replaced by the estimated cycloplegic autorefraction SER output by Equation 2.
286 Tested on Irish data alone, axial length estimated from spectacle refraction performed similarly to
287 that estimated using actual cycloplegic autorefraction SER (the latter shown in Table 2; MAE: 0.34 vs
288 0.34mm, respectively, concordance correlation coefficient: 0.91 vs 0.91, respectively, and coefficient

289 of repeatability: 0.83 vs 0.84mm respectively) and was better at identifying long eyes compared to
290 using spectacle SER alone (AUC 0.92 vs 0.81, respectively; Supplementary Figure 3).

291 **DISCUSSION**

292 We investigated and compared four machine learning methods for estimating axial length from
293 ocular and demographic variables. When tested on myopic eyes, the newly developed models all
294 performed better than previously published algorithms for axial length estimation. Interestingly,
295 standard multiple linear regression was the superior model, performing best on cross-sectional and
296 longitudinal data and meeting pre-specified criteria relevant to the clinical monitoring of axial length
297 change.

298 Machine learning model performance typically decreases when applied to new populations, known
299 as data or dataset shift^{34,35}; however, in this study we found that training region effects on
300 performance did not vary with testing region and that models were therefore applicable across
301 populations. This is, perhaps, unsurprising as axial length estimators are modelling optical
302 relationships, which are independent of populations. The reason why models tested on Chinese data
303 performed better is unclear, but could be related to differences in the age, SER or crystalline lens
304 power distributions.³⁶ Models trained using participants of only one sex (e.g. females) had a high
305 mean bias and a higher MAE when applied to the opposite sex (e.g. males), illustrating the
306 importance of incorporating sex into axial length estimation models. The linear regression algorithm
307 became slightly less precise with increasing age and high myopia.

308 Compared to previously published algorithms, the enhanced performance of the machine learning
309 models in this study is likely the result of the additional inclusion of age and sex as input variables
310 and the large, diverse training data set. Interestingly, the multiple linear regression model
311 outperformed the more complex deep learning models, indicating this is an effective method for
312 developing axial length estimation models. Adding ACD to models improved the estimation of axial
313 length, reducing the coefficient of repeatability by 0.08mm. While not commonly measured in

314 primary and community eye care settings, methods to non-biometrically measure ACD exist³⁷ and
315 may become more accessible in the future. Adding lens thickness will further improve axial length
316 estimation; however, clinicians are unlikely to have access to measures of lens thickness, but not
317 axial length. Other input variables could potentially be added to improve axial length estimation, for
318 example, axial length can be estimated from fundus photographs,^{38,39} but the performance of these
319 convolutional neural networks (MAE=0.74-0.90mm) was much poorer than models in the present
320 study (MAE=0.32 to 0.41mm).

321 Interestingly, adding ACD did not improve the estimation of axial length change (Figure 3). This is
322 likely because the axial length estimation errors, which largely arise from the ACD and crystalline
323 lens, are highly correlated at baseline and 12 months (Pearson $r=0.90$ in testing data). Thus,
324 subtracting the estimated axial length at baseline from 12 months cancels out the ACD contribution
325 to estimation error at these two visits.

326 The optimal linear regression model had a mean bias close to zero and a coefficient of repeatability
327 of ± 0.80 mm. While this coefficient of repeatability interval may seem large, it is approaching the
328 physiological limits of axial length prediction (± 0.75 mm), without incorporating effective lens power.
329 The ability of the linear regression model to identify a long eye (axial length in the top 2% for age
330 and sex) was much better than SER alone, representing a clinically useful application for identifying
331 eyes at higher risk of future vision loss. Likewise, the linear regression model could estimate the axial
332 length change over 12-months; however, it was not better than SER at identifying fast axial eye
333 growth. Previous authors have noted the difficulty in using axial length 'calculators' to obtain
334 appropriately sensitive measures of axial length change.¹³

335 Strengths of this study include the large sample sizes, comprehensive assessment of multiple
336 machine learning models and input variables, a focus on the clinical application of the tool and the
337 in-depth exploration of the expected limits of axial length estimation. Limitations of this study
338 generally relate to the application of its findings. Although models were trained using all refractive

339 errors, axial length estimation was only validated for myopic eyes, and our results may be less
340 applicable to non-myopic eyes. This could also be considered a strength given the likely clinical
341 application of an axial length estimator. Our results are similarly limited to ages 6-20 years and
342 model performance was less precise among highly myopic, older eyes. Extending the top-performing
343 linear regression to older ages generated implausible values (Supplementary Figure 4). Thus, further
344 work is needed to extend axial length estimation to older ages, but we have laid groundwork for
345 such an extension. To maximise data availability, we used keratometry measured by optical
346 biometers, rather than from autorefractors; however, strong agreement between these devices
347 means our results should be valid across both.^{40,41} The algorithm also requires cycloplegic SER, which
348 can be time-consuming and invasive to measure; however, we showed that extending the model to
349 use spectacle refraction, by adding an intermediate conversion step, produced good results for axial
350 length estimation among the Irish testing data and was better than spectacle SER alone for
351 identifying eyes with a long axial length. Such an intermediate step might also work for estimating
352 axial length from non-cycloplegic autorefraction⁴²; although, the increased variance of the internal
353 refractive power contribution to SER means axial length estimation would theoretically be less
354 precise than with cycloplegic autorefraction.

355 *Conclusion*

356 This study demonstrated that the inclusion of age and sex significantly improves axial length
357 estimation, and accuracy is further enhanced by adding anterior chamber depth. Multiple linear
358 regression outperformed more complicated machine learning models, was applicable across both
359 populations, was better than SER at identifying myopic eyes with long, high-risk axial length and
360 approached the physiological limits of axial length estimation without incorporating lens power. In
361 the absence of a biometer, the estimation of axial length from ocular and demographic variables
362 provides a useful tool in enabling the clinical management of myopia.

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364 **COMPETING INTERESTS**

365 GL and DSP are employees of Ocumetra, and JL and DIF are co-founders of Ocumetra, a company
366 providing data analytic tools to assist with the clinical management, including an axial length
367 estimation tool. JL is a consultant/contractor for Dopavision, Topcon, EssilorLuxottica and Ebiga
368 Vision and has received funding from Topcon, Ocumension, Kubota Vision, EssilorLuxottica, Vyluma,
369 Dopavision and Coopervision, all in the area of myopia management. DIF is a consultant/contractor
370 for Vyluma, Coopervision, Essilor, Thea, Ocumension and Johnson & Johnson and has received
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372 of research funding from HOYA Vision and Vyluma in the area of myopia management.

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374 **DATA AVAILABILITY:** The datasets generated during and/or analysed during the current study are
375 available from the corresponding author on reasonable request.

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480 **Figure 1** The performance of the four newly developed machine learning models (linear regression,
481 gradient boosting, multilayer perceptron, symbolic regression; shown in colour), as well as published
482 axial length prediction algorithms (no colour). Absolute mean bias is the absolute of the mean
483 difference between estimated and actual axial length. Better performance is indicated by a higher
484 concordance correlation coefficient and a lower value for all other metrics shown. Due to the high
485 correlation between the mean absolute error and the coefficient of repeatability (Pearson
486 correlation=0.99), only the latter is presented. Models were tested on either cross-sectional axial
487 length data (top; combined Irish and Chinese data) or longitudinal 12-month change in axial length
488 data (bottom; Irish data only).

489 **Figure 2** Relationships between predicted axial length and age in the four machine learning models.
490 The top row shows the predicted axial length at each age for a female with corneal radius of
491 curvature of 7.76mm and with spherical equivalent refraction (SER) of +3 dioptres (D), plano, -3 D
492 and -6 D. The bottom row shows the rate of change (1st derivative) of the plano curve (blue points)
493 compared to a best-fit spline curve (green line). The sum of the distances of all points from the best-
494 fit line are used as a measure of smoothness (uses +3 D, plano, -3 D and -6 D curves; linear
495 regression=0.19, gradient boosting=1.78, multilayer perceptron=0.19, symbolic regression=0.20). An
496 age curve is monotonic when all points for all age curves (+3 D, plano, -3 D and -6 D) lie on or above
497 0 (red dashed line). All models were monotonic except the multilayer perceptron model, which had
498 12.5% of first derivative points below 0.

499 **Figure 3** Performance of linear regression models as assessed by (a) agreement between predicted
500 and actual axial length at baseline using a Bland-Altman plot with mean bias and 95% limits of
501 agreement (LOA) shown; (b), ability of models to identify participants with axial length in the top 2%
502 for age and sex using a receiver operating characteristic (ROC) and area under the curve (AUC;
503 AXLest is the AUC for the machine learning model, SER is using spherical equivalent refractive error
504 alone); (c), agreement between predicted and 12-month change in axial length of myopic eyes (Irish

505 data only) using Bland-Altman plot and statistics; (d), ability of models to identify participants (Irish
506 data only) with 12-month axial length change greater than 0.2mm/year using a receiver operating
507 characteristic (ROC) and area under the curve. Linear regression models either (left column) do not
508 use anterior chamber depth (ACD; middle column), ACD as estimated from age and spherical
509 equivalent or (right column) actual ACD measured by biometry.

510 **Supplementary Figure 1** Data on 98th axial length centile thresholds extracted from the LIFE child
511 study and used to define eyes with a high-risk axial length. Points show the actual reported data
512 points and the lines show the smoothing cubic spline fit.

513 **Supplementary Figure 2** Scatter plot of axial length estimation error over age (left) and spherical
514 equivalent (right). Quantile regression slopes are shown for the 2.5% and 97.5% (dashed line), 25%
515 and 75% (dotted line) and 50% quantiles (solid line). Slopes are derived from multivariable quantile
516 regression including age and sex as explanatory variables. Slope positions on the left plot are
517 calculated assuming spherical equivalent of -3.12 dioptres (D) and slopes on the right are plotted
518 assuming an age of 12.5 years (both mean values of the testing data set).

519 **Supplementary Figure 3** Performance of the top-performing linear regression model when applied
520 to spectacle refraction data using Equation 2, using (top) a Bland-Altman plot with mean bias and
521 95% limits of agreement (LOA) shown, and (bottom), ability of models to identify participants with
522 axial length in the top 2% for age and sex using a receiver operating characteristic (ROC) and area
523 under the curve (AUC; AXLest is the AUC for the machine learning model, SER is using spherical
524 equivalent refractive error alone). The left-hand column is cross-sectional data, while the right-hand
525 column is longitudinal data. Testing data in all plots is from the Irish cohort alone.

526 **Supplementary Figure 4** Examples of the estimated axial length vs age function for females when
527 spherical equivalent and corneal radius are held constant at 0 D and 7.76mm, respectively, within
528 the age range investigated in the present study (left) and for ages up to 80 years (right)